# Introduction to multiple regression 

## Multiple regression

- Simple linear regression: Bivariate - two variables: $y$ and $x$
- Multiple linear regression: Multiple variables: $y$ and $x_{1}, x_{2}, \ldots$


## Weights of books

|  | weight $(\mathrm{g})$ | volume $\left(\mathrm{cm}^{3}\right)$ | cover |
| ---: | ---: | ---: | :---: |
| 1 | 800 | 885 | hc |
| 2 | 950 | 1016 | hc |
| 3 | 1050 | 1125 | hc |
| 4 | 350 | 239 | hc |
| 5 | 750 | 701 | hc |
| 6 | 600 | 641 | hc |
| 7 | 1075 | 1228 | hc |
| 8 | 250 | 412 | pb |
| 9 | 700 | 953 | pb |
| 10 | 650 | 929 | pb |
| 11 | 975 | 1492 | pb |
| 12 | 350 | 419 | pb |
| 13 | 950 | 1010 | pb |
| 14 | 425 | 595 | pb |
| 15 | 725 | 1034 | pb |



## Weights of books (cont.)

The scatterplot shows the relationship between weights and volumes of books as well as the regression output.


Books that are $10 \mathrm{~cm}^{3}$ over average are expected to weigh 7 g over average.

## Modeling weights of books using volume

## somewhat abbreviated output...

Coefficients:

|  | Estimate | Std. Error t value | Pr $(>\|t\|)$ |  |
| :--- | ---: | ---: | ---: | ---: |
| (Intercept) | 107.67931 | 88.37758 | 1.218 | 0.245 |
| Volume | 0.70864 | 0.09746 | 7.271 | $6.26 e-06$ |

Residual standard error: 123.9 on 13 degrees of freedom Multiple R-squared: 0.8026,Adjusted R-squared: 0.7875
F-statistic: 52.87 on 1 and 13 DF, p-value: 6.262e-06

## Weights of hardcover and paperback books

Can you identify a trend in the relationship between volume and weight of hardcover and paperback books?


## Weights of hardcover and paperback books

Can you identify a trend in the relationship between volume and weight of hardcover and paperback books?
Paperbacks generally weigh less than hardcover books after controlling for the book's volume.


## Qualitative predictors

## How can we include hardcover/paperback in our regression?

When a predictor takes two (categorical) values, we create a dummy variable which takes the values 0/1:

$$
\text { cover }: p b=\left\{\begin{array}{l}
0, \text { if cover is hardcover } \\
1, \text { if cover is paperback }
\end{array}\right.
$$

Which value we select for 0 is arbitrary, and is called the reference value

## Qualitative predictors

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1, \text { if cover is paperback }
\end{array}\right.
$$

How can we interpret $\beta_{1}$ in this case
Assume $y=\beta_{0}+\beta_{1} x+\epsilon_{i}$, where $y$ is weight, $x$ is cover: $p b$ Then

$$
y_{i}=\left\{\begin{array}{l}
\beta_{0}+\epsilon_{i}, \quad \text { if the cover is hardcover } \\
\beta_{0}+\beta_{1}+\epsilon_{i}, \quad \text { if the cover is paperback }
\end{array}\right.
$$

Paperback books are on average $\beta_{1}$ units heavier than hardcover books Do we expect $\beta_{1}$ to be positive or negative?

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\beta_{0}+\beta_{1}+\epsilon_{i}, \quad \text { if the cover is paperback }
\end{array}\right.
$$

Paperback books are on average $\beta_{1}$ units heavier than hardcover books Do we expect $\beta_{1}$ to be positive or negative?

## Qualitative predictors

When a predictor takes three (categorical) values (e.g., religion: Muslim/Christian/Atheist) we create two dummy variables

$$
\begin{aligned}
\text { rel: } \text { muslim } & =\left\{\begin{array}{l}
0, \text { if subject is NOT a Muslim } \\
1, \text { if subject is a Muslim }
\end{array}\right. \\
\text { rel: christian } & =\left\{\begin{array}{l}
0, \text { if subject is NOT a Christian } \\
1, \text { if subject is a Christian }
\end{array}\right.
\end{aligned}
$$

Why not three?

## Modeling weights of books using volume and cover type

Coefficients:

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|t\|)$ |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 197.96284 | 59.19274 | 3.344 | 0.005841 | ** |
| volume | 0.71795 | 0.06153 | 11.669 | $6.6 e-08$ | *** |
| cover:pb | -184.04727 | 40.49420 | -4.545 | 0.000672 | *** |

Residual standard error: 78.2 on 12 degrees of freedom
Multiple R-squared: 0.9275 ,Adjusted R-squared: 0.9154 F -
statistic: 76.73 on 2 and 12 DF, $\quad$-value: $1.455 e-$
07

## Visualising the linear model



Still a least squares solution.

## Modeling conditions

$$
\hat{y}=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\cdots+\beta_{p} x_{p}
$$

The model depends on the following conditions

1. residuals are nearly normal (less important for larger data sets)
2. residuals have constant variance
3. residuals are independent
4. each variable is linearly related to the outcome

## Determining the reference level

Based on the regression output below, which level of cover is the reference level? Note that pb: paperback.

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 197.9628 | 59.1927 | 3.34 | 0.0058 |
| volume | 0.7180 | 0.0615 | 11.67 | 0.0000 |
| cover:pb | -184.0473 | 40.4942 | -4.55 | 0.0007 |

a. paperback
b. hardcover

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a. paperback
b. hardcover

## Determining the reference level

Which of the below correctly describes the roles of variables in this regression model?

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
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a. response: weight, explanatory: volume, paperback cover
b. response: weight, explanatory: volume, hardcover cover
c. response: volume, explanatory: weight, cover type
d. response: weight, explanatory: volume, cover type

## Determining the reference level

Which of the below correctly describes the roles of variables in this regression model?

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c. response: volume, explanatory: weight, cover type
d. response: weight, explanatory: volume, cover type

## Linear Model

| Estimate |  |  | Std. Error | t value |
| :---: | :---: | :---: | :---: | :---: |
| (Intercept) | 197.96 | 59.19 | 3.34 | 0.01 |
| volume | 0.72 | 0.06 | 11.67 | 0.00 |
| cover:pb | -184.05 | 40.49 | -4.55 | 0.00 |

$$
\widehat{\text { weight }}=197.96+0.72 \text { volume }-184.05 \text { cover }: p b
$$

1. For hardcover books: plug in 0 for cover

$$
\begin{aligned}
\widehat{\text { weight }} & =197.96+0.72 \text { volume }-184.05 \times 0 \\
& =197.96+0.72 \text { volume }
\end{aligned}
$$

2. For paperback books: plug in 1 for cover

$$
\begin{aligned}
\widehat{\text { weight }} & =197.96+0.72 \text { volume }-184.05 \times 1 \\
& =13.91+0.72 \text { volume }
\end{aligned}
$$

## Visualising the linear model



## Linear model

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 197.96 | 59.19 | 3.34 | 0.01 |
| volume | 0.72 | 0.06 | 11.67 | 0.00 |
| cover:pb | -184.05 | 40.49 | -4.55 | 0.00 |

Slope of volume: All else held constant, books that are 1 more cubic centimeter in volume tend to weigh about 0.72 grams more.

Slope of cover: All else held constant, the model predicts that paperback books weigh 184 grams lower than hardcover books.

Intercept: Hardcover books with no volume are expected on average to weigh 198 grams.

- Obviously, the intercept does not make sense in context. It only serves to adjust the height of the line.


## Prediction

Which of the following is the correct calculation for the predicted weight of a paperback book that is $600 \mathrm{~cm}^{3}$ ?

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 197.96 | 59.19 | 3.34 | 0.01 |
| volume | 0.72 | 0.06 | 11.67 | 0.00 |
| cover:pb | -184.05 | 40.49 | -4.55 | 0.00 |

(a) $197.96+0.72$ * $600-184.05$ * 1
(b) $184.05+0.72$ * $600-197.96$ * 1
(c) $197.96+0.72$ * $600-184.05$ * 0
(d) $197.96+0.72$ * $1-184.05$ * 600

## Prediction

Which of the following is the correct calculation for the predicted weight of a paperback book that is $600 \mathrm{~cm}^{3}$ ?

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 197.96 | 59.19 | 3.34 | 0.01 |
| volume | 0.72 | 0.06 | 11.67 | 0.00 |
| cover:pb | -184.05 | 40.49 | -4.55 | 0.00 |

(a) $197.96+0.72$ * $600-184.05$ * 1
(b) $184.05+0.72$ * $600-197.96$ * 1
(c) $197.96+0.72$ * $600-184.05$ * 0
(d) $197.96+0.72$ * $1-184.05$ * 600

## Another Example: Predicting Poverty

Response variable: Percentage of residents living in poverty

Explanatory variables:
Percentage of residents living in a metropolitan area
Percentage of residents that are white
Percentage of residents that are high-school graduates
Percentage of residents that live in a single-parent, female-led household

## Revisit: Modeling poverty



## Another look at $R^{2}$

$R^{2}$ can be calculated in three ways:
1.square the correlation coefficient of $x$ and $y$ (how we have been calculating it)
2.square the correlation coefficient of $y$ and $\hat{y}$
3. based on definition: $\quad R^{2}=\frac{\text { explained variability in } y}{\text { total variability in } y}$

$$
R^{2}=1-\frac{S S_{R E G}}{S S_{T O T}}
$$

## Sum of squares

Sum of squares of $y$ : $S S_{\text {тот }}=\sum_{i}(y-\bar{y})^{2}=480.25$
Sum of squares of residuals : $S S_{R E G}=\sum_{i}(y-\hat{y})^{2}=347.68$

$$
R^{2}=1-\frac{S S_{\text {REG }}}{S S_{T O T}}=0.29
$$

## R squared

- For single-predictor linear regression, this is the square of the calculation coefficient.
- However, in multiple linear regression, we can't calculate $R^{2}$ as the square of the correlation between $x$ and $y$ because we have multiple $x s$.
- And next we'll learn another measure of explained variability, adjusted $R^{2}$, that requires the use of the third approach, ratio of explained and unexplained variability.


## Adjusted $R^{2}$

|  | $R^{2}$ |
| :--- | :---: |
| Model 1 (Single-predictor) | 0.28 |
| Model 2 (Multiple) | 0.29 |

When any variable is added to the model $\mathrm{R}^{2}$ increases.
But it may be the case that the variable is not really informative (at all or in the context of the other variables)

This is called overfitting and can lead to falsely large $R^{2}$

|  | $R^{2}$ | Adjusted $R^{2}$ |
| :--- | :---: | :---: |
| Model 1 (Single-predictor) | 0.28 | 0.26 |
| Model 2 (Multiple) | 0.29 | 0.26 |

When any variable is added to the model $\mathrm{R}^{2}$ increases.
But if the added variable doesn't really provide any new information, or is completely unrelated, adjusted $\mathrm{R}^{2}$ does not increase.

## $R^{2}$ vs. adjusted $R^{2}$

## Adjusted $R^{2}$

$$
R_{a d j}^{2}=1-\left(\frac{S S_{R E G}}{S S_{T O T}} \times \frac{n-1}{n-p-1}\right)
$$

where $n$ is the number of cases and $p$ is the number of predictors (explanatory variables) in the model.

- Because $p$ is never negative, $R_{a d j}^{2}$ will always be smaller than $R^{2}$
- $R_{a d j}^{2}$ applies a penalty for the number of predictors included in the model.
- Therefore, we choose models with higher $R_{a d j}^{2}$ over others.


## Calculate adjusted $R^{2}$

$$
\begin{aligned}
S S_{\text {REG }} & =339.47 \\
S S_{T O T} & =480.25 \\
n & =51 \\
p & =2
\end{aligned}
$$

$$
\begin{aligned}
R_{a d j}^{2} & =1-\left(\frac{S S_{R E G}}{S S_{T O T}} \times \frac{n-1}{n-p-1}\right)= \\
\quad= & 1-\left(\frac{339.47}{480.25} \times \frac{51-1}{51-2-1}\right) \\
& =1-\left(\frac{339.47}{480.25} \times \frac{50}{48}\right) \\
& =1-0.74 \\
& =0.26
\end{aligned}
$$

Collinearity
Predicting poverty in US states: pair plot


## Predicting poverty using \% female hh + \% white

| Linear model: | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | -2.58 | 5.78 | -0.45 | 0.66 |
| female_house | 0.89 | 0.24 | 3.67 | 0.00 |

$$
R^{2}=1-\frac{S S_{R E G}}{S S_{T O T}}=0.29
$$

## Predicting poverty using \% female hh + \% white

| Linear model: | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | -2.58 | 5.78 | -0.45 | 0.66 |
| female_house | 0.89 | 0.24 | 3.67 | 0.00 |
| white | 0.04 | 0.04 | 1.08 | 0.29 |

$$
R^{2}=1-\frac{S S_{R E G}}{S S_{T O T}}=0.29
$$

Does adding the variable white to the model add valuable information that wasn't provided by female house?


## Collinearity between explanatory variables

Two predictor variables are said to be collinear when they are correlated, and this collinearity complicates model estimation.
Remember: Predictors are also called explanatory or independent variables. Ideally, they would be independent of each other.

## Collinearity between explanatory variables (cont.)

Two predictor variables are said to be collinear when they are correlated, and this collinearity complicates model estimation.
Remember: Predictors are also called explanatory or independent variables. Ideally, they would be independent of each other.

We don't like adding predictors that are associated with each other to the model, because often times the addition of such variable brings nothing to the table and may lead to overfitting. Instead, we prefer the simplest best model.

When we have multiple variables, we want to select a subset of these variables to include in the model.

## Beauty in the classroom

- Data: Student evaluations of instructors' beauty and teaching quality for 463 courses at the University of Texas.
- Evaluations conducted at the end of semester, and the beauty judgements were made later, by six students who had not attended the classes and were not aware of the course evaluations (2 upper level females, 2 upper level males, one lower level female, one lower level male).

Hamermesh \& Parker. (2004)"Beauty in the classroom: instructors' pulchritude and putative pedagogical productivity

Economics Education Review.

## Beauty in the classroom

> beauty
> gender.male
> age
> formal.yes ${ }^{1}$
> lower.yes $^{2}$
> native.non english
> minority.yes
> students $^{3}$
> tenure.tenure track ${ }^{4}$
> tenure.tenured

[^0]
## Professor rating vs. beauty

Professor evaluation score (higher score means better) vs. beauty score (a score of 0 means average, negative score means below average, and a positive score above average):


Which of the below is correct based on the model output?

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 4.19 | 0.03 | 167.24 | 0.00 |
| beauty | 0.13 | 0.03 | 4.00 | 0.00 |
| $R^{2}=0.0336$ |  |  |  |  |

(a) Model predicts 3.36\% of professor ratings correctly.
(b) Beauty is not a significant predictor of professor evaluation.
(c) Professors who score 1 point above average in their beauty score are tend to also score 0.13 points higher in their evaluation.
(d) 3.36\% of variability in beauty scores can be explained by professor evaluation.
(e) The correlation coefficient could be $\sqrt{0.0336}=0.18$ or -0.18 , we can't tell which is correct.

Which of the below is correct based on the model output?

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 4.19 | 0.03 | 167.24 | 0.00 |
| beauty | 0.13 | 0.03 | 4.00 | 0.00 |
| $R^{2}=0.0336$ |  |  |  |  |

(a) Model predicts 3.36\% of professor ratings correctly.
(b) Beauty is not a significant predictor of professor evaluation.
(c) Professors who score 1 point above average in their beauty score are tend to also score 0.13 points higher in their evaluation.
(d) $3.36 \%$ of variability in beauty scores can be explained by professor evaluation.
(e) The correlation coefficient could be $\sqrt{0.0336}=0.18$ or -0.18 , we can't tell which is correct.

## Exploratory analysis

Any interesting features?

For a given beauty score, are male professors evaluated higher, lower, or about the same as female professors?


## Professor rating vs. beauty + gender

For a given beauty score, are male professors evaluated higher, lower, or about the same as female professors?

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 4.09 | 0.04 | 107.85 | 0.00 |
| beauty | 0.14 | 0.03 | 4.44 | 0.00 |
| gender.male | 0.17 | 0.05 | 3.38 | 0.00 |
| $R_{a d j}^{2}=0.057$ |  |  |  |  |

(a) higher
(b) lower
(c) about the same

## Professor rating vs. beauty + gender

For a given beauty score, are male professors evaluated higher, lower, or about the same as female professors?

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 4.09 | 0.04 | 107.85 | 0.00 |
| beauty | 0.14 | 0.03 | 4.44 | 0.00 |
| gender.male | 0.17 | 0.05 | 3.38 | 0.00 |
| $R_{a d j}^{2}=0.057$ |  |  |  |  |

(a) higher $\rightarrow$ Beauty held constant, male professors are rated 0.17 points higher on average than female professors.
(b) lower
(c) about the same

## Full Model

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|t\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 4.6282 | 0.1720 | 26.90 | 0.00 |
| beauty | 0.1080 | 0.0329 | 3.28 | 0.00 |
| gender.male | 0.2040 | 0.0528 | 3.87 | 0.00 |
| age | -0.0089 | 0.0032 | -2.75 | 0.01 |
| formal.yes $^{1}$ | 0.1511 | 0.0749 | 2.02 | 0.04 |
| lower.yes $^{2}$ | 0.0582 | 0.0553 | 1.05 | 0.29 |
| native.non english $^{2}$ | -0.2158 | 0.1147 | -1.88 | 0.06 |
| minority.yes $^{3}$ | -0.0707 | 0.0763 | -0.93 | 0.35 |
| students $^{3}$ | -0.0004 | 0.0004 | -1.03 | 0.30 |
| tenure.tenure track $^{4}$ | -0.1933 | 0.0847 | -2.28 | 0.02 |
| tenure.tenured $^{2}$ | -0.1574 | 0.0656 | -2.40 | 0.02 |

[^1]
## Hypotheses

Just as the interpretation of the slope parameters take into account all other variables in the model, the hypotheses for testing for significance of a predictor also takes into account all other variables.
$H_{0}^{i}: \beta_{i}=0$ when other explanatory variables are included in the model. $H_{A}^{i}: \beta_{i} \neq 0$ when other explanatory variables are included in the model.

## Assessing significance: numerical variables

The $p$-value for age is 0.01 . What does this indicate?

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|t\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| $\ldots$ |  |  |  |  |
| age | -0.0089 | 0.0032 | -2.75 | 0.01 |
| $\ldots$ |  |  |  |  |

a. Since $p$-value is positive, higher the professor's age, the higher we would expect them to be rated.
b. If we keep all other variables in the model, there is strong evidence that professor's age is associated with their rating.
c. Probability that the true slope parameter for age is 0 is 0.01 .
d. There is about $1 \%$ chance that the true slope parameter for age is 0.0089.

## Assessing significance: numerical variables

The $p$-value for age is 0.01 . What does this indicate?

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| $\ldots$ <br> age | -0.0089 | 0.0032 | -2.75 | 0.01 |
| $\ldots$ |  |  |  |  |

a. Since $p$-value is positive, higher the professor's age, the higher we would expect them to be rated.
b. If we keep all other variables in the model, there is strong evidence that professor's age is associated with their rating.
c. Probability that the true slope parameter for age is 0 is 0.01 .
d. There is about $1 \%$ chance that the true slope parameter for age is 0.0089.

## Assessing significance

Which predictors do not seem to meaningfully contribute to the model, i.e. may not be significant predictors of professor's rating score?

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|t\|)$ |
| ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 4.6282 | 0.1720 | 26.90 | 0.00 |
| beauty | 0.1080 | 0.0329 | 3.28 | 0.00 |
| gender.male | 0.2040 | 0.0528 | 3.87 | 0.00 |
| age | -0.0089 | 0.0032 | -2.75 | 0.01 |
| formal.yes | 0.1511 | 0.0749 | 2.02 | 0.04 |
| lower.yes | 0.0582 | 0.0553 | 1.05 | 0.29 |
| native.non english | -0.2158 | 0.1147 | -1.88 | 0.06 |
| minority.yes | -0.0707 | 0.0763 | -0.93 | 0.35 |
| students | -0.0004 | 0.0004 | -1.03 | 0.30 |
| tenure.tenure track | -0.1933 | 0.0847 | -2.28 | 0.02 |
| tenure.tenured | -0.1574 | 0.0656 | -2.40 | 0.02 |

## Model selection strategies

Based on what we've learned so far, what are some ways you can think of that can be used to determine which variables to keep in the model and which to leave out?

## Backward-elimination

1. Start with the full model
2. Drop one variable at a time and record $R_{a d j}^{2}$ of each smaller model
3. Pick the model with the highest increase in $R_{a d j}^{2}$
4. Repeat until none of the models yield an increase in $R_{a d j}^{2}$

## Backward-elimination

| Full | beauty + gender + age + formal + lower + native + minority + students + tenure | 0.0839 |
| :---: | :---: | :---: |
| Step 1 | ```gender + age + formal + lower + native + minority + students + tenure beauty + age + formal + lower + native + minority + students + tenure beauty + gender + formal + lower + native + minority + students + tenure beauty + gender + age + lower + native + minority + students + tenure beauty + gender + age + formal + native + minority + students + tenure beauty + gender + age + formal + lower + minority + students + tenure beauty + gender + age + formal + lower + native + students + tenure beauty + gender + age + formal + lower + native + minority + tenure beauty + gender + age + formal + lower + native + minority + students``` | $\begin{aligned} & 0.0642 \\ & 0.0557 \\ & 0.0706 \\ & 0.0777 \\ & 0.0837 \\ & 0.0788 \\ & 0.0842 \\ & 0.0838 \\ & 0.0733 \end{aligned}$ |
| Step 2 | ```gender + age + formal + lower + native + students + tenure beauty + age + formal + lower + native + students + tenure beauty + gender + formal + lower + native + students + tenure beauty + gender + age + lower + native + students + tenure beauty + gender + age + formal + native + students + tenure beauty + gender + age + formal + lower + native + tenure beauty + gender + age + formal + lower + native + students``` | $\begin{aligned} & 0.0647 \\ & 0.0543 \\ & 0.0708 \\ & 0.0776 \\ & 0.0846 \\ & 0.0844 \\ & 0.0725 \end{aligned}$ |
| Step 3 | ```gender + age + formal + native + students + tenure beauty + age + formal + native + students + tenure beauty + gender + formal + native + students + tenure beauty + gender + age + native + students + tenure beauty + gender + age + formal + students + tenure beauty + gender + age + formal + native + tenure beauty + gender + age + formal + native + students``` | $\begin{aligned} & \hline 0.0653 \\ & 0.0534 \\ & 0.0707 \\ & 0.0786 \\ & 0.0756 \\ & 0.0855 \\ & 0.0713 \\ & \hline \end{aligned}$ |
| Step 4 | ```gender + age + formal + native + tenure beauty + age + formal + native + tenure beauty + gender + formal + native + tenure beauty + gender + age + native + tenure beauty + gender + age + formal + tenure beauty + gender + age + formal + native``` | $\begin{aligned} & \hline 0.0667 \\ & 0.0553 \\ & 0.0723 \\ & 0.0806 \\ & 0.0773 \\ & 0.0713 \end{aligned}$ |

## Forward-selection

1. Start with regressions of response vs. each explanatory variable
2. Pick the model with the highest $R_{a d j}^{2}$
3. Add the remaining variables one at a time to the existing model, and once again pick the model with the highest $R_{a d j}^{2}$
4. Repeat until the addition of any of the remanning variables does not result in a higher $R_{a d j}^{2}$

## Backward-Elimination vs. Forward-Selection

Backward elimination with the $p$-value approach:

1. Start with the full model
2. Drop the variable with the highest $p$-value and refit a smaller model
3. Repeat until all variables left in the model are significant

Forward selection with the p-value approach:

1. Start with regressions of response vs. each explanatory variable
2. Pick the variable with the lowest significant $p$-value
3. Add the remaining variables one at a time to the existing model, and pick the variable with the lowest significant $p$-value
4. Repeat until any of the remaining variables does not have a significant p -value

## Adjusted $R^{2}$ vs. p-value approaches

- The two approaches are similar, but they sometimes lead to different models, with the adjusted $R^{2}$ approach tending to include more predictors in the final model.
- When the sole goal is to improve prediction accuracy, use $R^{2}$. This is commonly the case in machine learning applications.
- When we care about understanding which variables are statistically significant predictors of the response, or if there is interest in producing a simpler model at the potential cost of a little prediction accuracy, then the $p$-value approach is preferred.
- Regardless of the approach we use, our job is not done after variable selection - we must still verify the model conditions are reasonable.


## Checking model conditions using graphs

## Modeling conditions

$$
\hat{y}=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\cdots+\beta_{p} x_{p}
$$

The model depends on the following conditions

1. residuals are nearly normal (less important for larger data sets)
2. residuals have constant variability
3. residuals are independent
4. each variable is linearly related to the outcome

We often use graphical methods to check the validity of these conditions, which we will go through in detail in the following slides.

## (1) nearly normal residuals

Histogram of the residuals.
Histogram of residuals


Does this condition appear to be satisfied?

## (2) constant variability in residuals

Scatterplot of residuals and/or absolute value of residuals vs. fitted (predicted).



Does this condition appear to be satisfied?

## Checking constant variance - recap

When we did simple linear regression (one explanatory variable) we checked the constant variance condition using a plot of residuals vs. $x$.

With multiple linear regression (2+ explanatory variables) we checked the constant variance condition using a plot of residuals vs. fitted.

Why are we using different plots?

## Checking constant variance - recap

When we did simple linear regression (one explanatory variable) we checked the constant variance condition using a plot of residuals vs. $x$.

With multiple linear regression (2+ explanatory variables) we checked the constant variance condition using a plot of residuals vs. fitted.

## Why are we using different plots?

In multiple linear regression there are many explanatory variables, so a plot of residuals vs. one of them wouldn't give us the complete picture.

## (4) linear relationships

Scatterplot of residuals vs. each (numerical) explanatory variable.


Does this condition appear to be satisfied?

## Several options for improving a model

## Transforming variables

Seeking out additional variables to fill model gaps

Using more advanced methods that would account for challenges around inconsistent variability or nonlinear relationships between predictors and the outcome

## Transformations

If the concern with the model is non-linear relationships between the explanatory variable(s) and the response variable, transforming the response variable can be helpful.

- Log transformation (log y)
- Square root transformation (sqrt(y))
- Inverse transformation (1/y)

It is also possible to apply transformations to the explanatory variable(s), however such transformations tend to make the model coefficients even harder to interpret.

## Models can be wrong, but useful

All models are wrong, but some are useful.

- George Box

No model is perfect, but even imperfect models can be useful, as long as we are clear and report the model's shortcomings.

If conditions are grossly violated, we should not report the model results, but instead consider a new model, even if it means learning more statistical methods or hiring someone who can help.


[^0]:    ${ }^{1}$ formal: picture wearing tie\&jacket/blouse, levels: yes, no
    ${ }^{2}$ lower: lower division course, levels: yes, no
    ${ }^{3}$ students: number of students
    ${ }^{4}$ tenure: tenure status, levels: non-tenure track, tenure track, tenured

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