## Useful Distributions (pt 2)

## Negative binomial distribution

## Practice

A college student working at a psychology lab is asked to recruit 10 couples to participate in a study.
She decides to stand outside the student center and ask every person leaving the building whether they are in a relationship and, if so, whether they would like to participate in the study with their significant other.
Suppose the probability of finding such a person is $10 \%$. What is the probability that she will need to ask 30 people before she hits her goal?

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## Negative binomial distribution

- The negative binomial distribution describes the probability of observing the k-th success on the nth trial.
- The following four conditions are useful for identifying a negative binomial case:

1. The trials are independent.
2. Each trial outcome can be classified as a success or failure.
3. The probability of success $(p)$ is the same for each trial.
4. The last trial must be a success.

Note that the first three conditions are common to the binomial distribution.

Negative binomial distribution

$$
\mathrm{P}\left(k^{t h} \text { success on the } n^{t h} \text { trial }\right)=\binom{n-1}{k-1} p^{k}(1-p)^{n-k}
$$

where p is the probability that an individual trial is a success. All trials are assumed to be independent.

## Negative binomial distribution

|  | $X$ is counting... | Probability mass function | Formula | Alternate formula (using equivalent binomial) | Alternate formula <br> (simplified using: $n=k+r)$ | Support |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | $k$ failures, given $r$ successes | $f(k ; r, p) \equiv \operatorname{Pr}(X=k)=$ | $\underset{\substack{k+r-1 \\[7][5][8]}}{(8]}) p^{r}(1-p)^{k}$ | $\left.{ }_{[2]}^{(2+r-1}\right) p^{r}(1-p)^{k}$ <br> [9][10][11] | $\binom{n-1}{k} p^{r}(1-p)^{k}$ | for $k=0,1,2, \ldots$ |
| 2 | $n$ trials, given <br> $r$ successes | $f(n ; r, p) \equiv \operatorname{Pr}(X=n)=$ | $\binom{n-1}{r-1} p^{r}(1-p)^{n-r}$ <br> [5][11][12][13][14] | $\binom{n-1}{n-r} p^{r}(1-p)^{n-r}$ |  | for $n=r, r+1, r+2, \ldots$ |
| 3 | $n$ trials, given $r$ failures | $f(n ; r, p) \equiv \operatorname{Pr}(X=n)=$ | $\binom{n-1}{r-1} p^{n-r}(1-p)^{r}$ | $\binom{n-1}{n-r} p^{n-r}(1-p)^{r}$ | $\binom{n-1}{k} p^{k}(1-p)^{r}$ |  |
| 4 | $r$ successes, given $n$ trials | $f(r ; n, p) \equiv \operatorname{Pr}(X=r)=$ | This is the binomial distribution: $\binom{n}{r} p^{r}(1-p)^{n-r}$ |  |  | for $r=0,1,2, \ldots, n$ |

## Negative binomial distribution

An RV $X$ has the Negative Binomial distribution with parameters $r$ and $p$ if it has the pf

$$
f(x \mid p, r)=\left\{\begin{array}{cc}
\binom{r+x-1}{x} p^{r}(1-p)^{x} & x=0,1,2, \ldots \\
0 & \text { otherwise }
\end{array}\right.
$$

where $0<p<1$ and $r$ is a positive integer.

Say we have an infinite sequence of Bernoulli trials with parameter $p$, and $X=$ number of "failures" before the $r$ th "success". Then $X \sim \operatorname{Neg} \operatorname{Binomial}(r, p)$.

- Parameter space: $r$ positive integer, $p \in(0,1)$.
- $E(X)=\frac{r(1-p)}{p} \quad \operatorname{Var}(X)=\frac{r(1-p)}{p^{2}}$


## Properties

Theorem (Sum of Geometric is Negative Binomial)

If $X_{1}, \ldots, X_{r}$ are i.i.d. and each $X_{i} \sim \operatorname{Geometric}(p)$ then $X$ $=X_{1}+\cdots+X_{r} \sim \operatorname{NegBinomial}(r, p)$.

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$P\left(10^{\text {th }}\right.$ success on the $30^{\text {th }}$ trial $)$

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## Binomial vs negative binomial

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How is the negative binomial distribution different from the binomial distribution?

- In the binomial case, we typically have a fixed number of trials and instead consider the number of successes.
- In the negative binomial case, we examine how many trials it takes to observe a fixed number of successes and require that the last observation be a success.


## Practice

Which of the following describes a case where we would use the negative binomial distribution to calculate the desired probability?
(a) Probability that a 5 year old boy is taller than 42 inches.
(b) Probability that 3 out of 10 softball throws are successful.
(c) Probability of being dealt a straight flush hand in poker.
(d) Probability of missing 8 shots before the first hit.
(e) Probability of hitting the ball for the 3rd time on the 8th try.

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## Poisson Distribution

## Poisson Distribution

Wight et al (2004) looked at the variation in cadaveric heart beating organ donor rates in the UK.
There were 1330 organ donors, aged 15-69, across the UK for the two years 1999 and 2000 combined.
$X$ : Number of donors per day
$P(X)$ ?

The mean number of organ donors per day over the two-year period is calculated as:
$\lambda=\frac{1330}{365+365}=\frac{1330}{730}=1.82$ organ donations per day

82 million person years, each person has a very small probability of becoming an organ donor.

## Poisson distribution

- The Poisson distribution is often useful for estimating the number of rare events in a large population over a short unit of time for a fixed population if the individuals within the population are independent.
- The rate for a Poisson distribution is the average number of occurrences in a mostly-fixed population per unit of time, and is typically denoted by $\lambda$.
- Using the rate, we can describe the probability of observing exactly $k$ events in a single unit of time.

$$
P(X=k)=\frac{\lambda^{k} e^{-\lambda}}{k!}, k=0,1,2, \ldots
$$

## Poisson distribution



## Poisson or Binomial

Let $X \sim \operatorname{Binomial}\left(n, p=\frac{\lambda}{n}\right)$, where $\lambda>0$ is fixed. Then for any $k \in\{0,1,2, \ldots\}$, we have

$$
\lim _{n \rightarrow \infty} P_{X}(k)=\frac{e^{-\lambda} \lambda^{k}}{k!} .
$$

Proof
We have

$$
\begin{gathered}
\lim _{n \rightarrow \infty} P_{X}(k)=\lim _{n \rightarrow \infty}\binom{n}{k}\left(\frac{\lambda}{n}\right)^{k}\left(1-\frac{\lambda}{n}\right)^{n-k} \\
=\lambda^{k} \lim _{n \rightarrow \infty} \frac{n!}{k!(n-k)!}\left(\frac{1}{n^{k}}\right)\left(1-\frac{\lambda}{n}\right)^{n-k} \\
=\frac{\lambda^{k}}{k!} \cdot \lim _{n \rightarrow \infty}\left(\left[\frac{n(n-1)(n-2) \ldots(n-k+1)}{n^{k}}\right]\left[\left(1-\frac{\lambda}{n}\right)^{n}\right]\left[\left(1-\frac{\lambda}{n}\right)^{-k}\right]\right) .
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1
\end{array} e^{-\lambda}
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P(\text { only } 1 \text { failure in a week }) & =\frac{2^{1} \times e^{-2}}{1!} \\
& =\frac{2 \times e^{-2}}{1} \\
& =0.27
\end{aligned}
$$

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We are given the weekly failure rate, but to answer this question we need to first calculate the average rate of failure on a given day: $\lambda_{\text {day }}=2 / 7=0.2857$. Note that we are assuming that the probability 7 of power failure is the same on any day of the week, i.e. we assume independence.

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P(3 \text { failures on a given day }) & =\frac{0.2857^{1} \times e^{-0.2857}}{3!} \\
& =\frac{0.2857 \times e^{-0.2857}}{6} \\
& =0.0358
\end{aligned}
$$

## Is it Poisson?

- A random variable may follow a Poisson distribution if the event being considered is rare, the population is large, and the events occur independently of each other
- However we can think of situations where the events are not really independent. For example, if we are interested in the probability of a certain number of weddings over one summer, we should take into consideration that weekends are more popular for weddings.
- In this case, a Poisson model may sometimes still be reasonable if we allow it to have a different rate for different times; we could model the rate as higher on weekends than on weekdays.
- The idea of modeling rates for a Poisson distribution against a second variable (day of the week) forms the foundation of some more advanced methods called generalized linear models.


## Exponential Distribution

Suppose that we just had an electricity power failure. What is the expected time until the next failure?

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Suppose that we just had an electricity power failure. What is the expected time until the next failure?

Let $\lambda>0$. A random variable $X$ follows the exponential distribution with parameter $\lambda$ if it has a continuous distribution with pf:

$$
f(x \mid \lambda)=\left\{\begin{array}{cc}
\lambda e^{-\lambda x} & x>0 \\
0 & \text { otherwise }
\end{array}\right.
$$

Parameter space: $\lambda \in[0, \infty)$.

$$
E(X)=\frac{1}{\lambda}, \quad \operatorname{Var}(X)=\frac{1}{\lambda^{2}}
$$

## Exponential Distribution

## Relationship of the Poisson and Exponential distribution

$X$ : Time between two consecutive arrivals
$N_{t}$ : Number of arrivals during time $t$

$$
\begin{gathered}
P(X>x)=P\left(N_{t}=N_{t+x}\right)=P\left(N_{x}=0\right)=\frac{\lambda^{0} e^{-\lambda}}{0!}=e^{-\lambda} \\
P(X \leq x)=1-P(X<x)=1-e^{-\lambda}
\end{gathered}
$$

## Memorylessness of the Exponential

$$
\begin{aligned}
& P(X>x+a \mid X>a)= \frac{P(X>x+a, X>a)}{P(X>a)} \\
&= \frac{P(X>x+a)}{P(X>a)} \\
&= \frac{1-F_{X}(x+a)}{1-F_{X}(a)} \\
&= \frac{e^{-\lambda(x+a)}}{e^{-\lambda a}} \\
&=e^{-\lambda x} \\
&=P(X>x) .
\end{aligned}
$$

## Poisson Example

The number of emails that I get in a weekday can be modelled by a Poisson distribution with an average of 0.2 emails per minute.

1. What is the probability that I get no emails in an interval of length 5 minutes?
2. What is the probability that I get more than 3 emails in an interval of length 10 minutes?
3. I just got an email. What is the probability that I will wait more than 3 minutes until the next email?
