Useful Distributions (pt 2)

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- The negative binomial distribution describes the probability of observing the k-th success on the nth trial.
- The following four conditions are useful for identifying a negative binomial case:
 - 1. The trials are independent.
 - 2. Each trial outcome can be classified as a success or failure.
 - 3. The probability of success (p) is the same for each trial.
 - 4. The last trial must be a success.

Note that the first three conditions are common to the binomial distribution.

Negative binomial distribution

$$P(k^{th} \text{ success on the } n^{th} \text{ trial}) = \binom{n-1}{k-1} p^k (1-p)^{n-k},$$

where p is the probability that an individual trial is a success. All trials are assumed to be independent.

	<i>X</i> is counting	Probability mass function	Formula	Alternate formula (using equivalent binomial)	Alternate formula $({f simplified\ using:}\ n=k+r)$	Support
1	<i>k</i> failures, given <i>r</i> successes	$f(k;r,p)\equiv \Pr(X=k)=$	${{k+r-1}\choose{k}}p^r(1-p)^k$ [7][5][8]	$\binom{k+r-1}{r-1}p^r(1-p)^k$ [2] [9][10][11]	$\binom{n-1}{k}p^r(1-p)^k$	$\text{for } k=0,1,2,\ldots$
2	<i>n</i> trials, given <i>r</i> successes	$f(n;r,p)\equiv \Pr(X=n)=$	$\binom{n-1}{r-1}p^r(1-p)^{n-r}$ [5][11][12][13][14]	${n-1 \choose n-r}p^r(1-p)^{n-r}$		for $n=r,r+1,r+2,\ldots$
3	<i>n</i> trials, given <i>r</i> failures	$f(n;r,p)\equiv \Pr(X=n)=$	${n-1 \choose r-1}p^{n-r}(1-p)^r$	${n-1 \choose n-r}p^{n-r}(1-p)^r$	$\binom{n-1}{k}p^k(1-p)^r$	
4	<i>r</i> successes, given <i>n</i> trials	$f(r;n,p)\equiv \Pr(X=r)=$	This is the binomial distribution: $\binom{n}{r}p^r(1-p)^{n-r}$			$\text{for } r=0,1,2,\ldots,n$

An RV *X* has the Negative Binomial distribution with parameters *r* and *p* if it has the pf

$$f(x \mid p, r) = \begin{cases} \binom{r+x-1}{x} p^r (1-p)^x & x = 0, 1, 2, \dots, \\ 0 & \text{otherwise.} \end{cases}$$

where 0 and <math>r is a positive integer.

Say we have an infinite sequence of Bernoulli trials with parameter p, and X = number of "failures" before the r th "success". Then $X \sim$ Neg Binomial (r, p).

• Parameter space: r positive integer, $p \in (0,1)$.

•
$$E(X) = \frac{r(1-p)}{p}$$
 $Var(X) = \frac{r(1-p)}{p^2}$

Properties

Theorem (Sum of Geometric is Negative Binomial)

If $X_1, ..., X_r$ are i.i.d. and each $X_i \sim \text{Geometric}(p)$ then $X = X_1 + \cdots + X_r \sim \text{NegBinomial}(r, p)$.

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Suppose the probability of finding such a person is 10%. What is the probability that she will need to ask 30 people before she hits her goal?

 $P(10^{th} \text{ success on the } 30^{th} \text{ trial})$

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$$= 0.00012$$

Binomial vs negative binomial

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How is the negative binomial distribution different from the binomial distribution?

- In the binomial case, we typically have a fixed number of trials and instead consider the number of successes.
- In the negative binomial case, we examine how many trials it takes to observe a fixed number of successes and require that the last observation be a success.

Which of the following describes a case where we would use the negative binomial distribution to calculate the desired probability?

(a) Probability that a 5 year old boy is taller than 42 inches.

- (b) Probability that 3 out of 10 softball throws are successful.
- (c) Probability of being dealt a straight flush hand in poker.
- (d) Probability of missing 8 shots before the first hit.
- (e) Probability of hitting the ball for the 3rd time on the 8th try.

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Poisson Distribution

Poisson Distribution

Wight et al (2004) looked at the variation in cadaveric heart beating organ donor rates in the UK.

There were 1330 organ donors, aged 15-69, across the UK for the two years 1999 and 2000 combined.

X: Number of donors per day P(X)?

The mean number of organ donors per day over the two-year period is calculated as:

$$\lambda = \frac{1330}{365+365} = \frac{1330}{730} = 1.82$$
 organ donations per day

82 million person years, each person has a very small probability of becoming an organ donor.

Poisson distribution

- The Poisson distribution is often useful for estimating the number of rare events in a large population over a short unit of time for a fixed population if the individuals within the population are independent.
- The rate for a Poisson distribution is the average number of occurrences in a mostly-fixed population per unit of time, and is typically denoted by λ.
- Using the rate, we can describe the probability of observing exactly *k* events in a single unit of time.

$$P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}, k = 0, 1, 2, ...$$

Poisson distribution



Poisson or Binomial

Let $X \sim \text{Binomial}\left(n, p = \frac{\lambda}{n}\right)$, where $\lambda > 0$ is fixed. Then for any $k \in \{0, 1, 2, ...\}$, we have

$$\lim_{n\to\infty}P_X(k)=\frac{e^{-\lambda}\lambda^k}{k!}.$$

Proof We have

$$\lim_{n \to \infty} P_X(k) = \lim_{n \to \infty} {\binom{n}{k}} \left(\frac{\lambda}{n}\right)^k \left(1 - \frac{\lambda}{n}\right)^{n-k}$$
$$= \lambda^k \lim_{n \to \infty} \frac{n!}{k! (n-k)!} \left(\frac{1}{n^k}\right) \left(1 - \frac{\lambda}{n}\right)^{n-k}$$
$$= \frac{\lambda^k}{k!} \cdot \lim_{n \to \infty} \left(\left[\frac{n(n-1)(n-2)\dots(n-k+1)}{n^k}\right] \left[\left(1 - \frac{\lambda}{n}\right)^n \right] \left[\left(1 - \frac{\lambda}{n}\right)^{-k} \right] \right).$$

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$$1 \qquad e^{-\lambda} \qquad 1$$

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$$P(\text{only 1 failure in a week}) = \frac{2^1 \times e^{-2}}{1!}$$
$$= \frac{2 \times e^{-2}}{1}$$
$$= 0.27$$

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We are given the weekly failure rate, but to answer this question we need to first calculate the average rate of failure on a given day: $\lambda_{day} = 2/7 = 0.2857$. Note that we are assuming that the probability 7 of power failure is the same on any day of the week, i.e. we assume independence.

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$$= \frac{0.2857^{1} \times e^{-0.2857}}{3!}$$
$$= \frac{0.2857 \times e^{-0.2857}}{6}$$
$$= 0.0358$$

Is it Poisson?

- A random variable may follow a Poisson distribution if the event being considered is rare, the population is large, and the events occur independently of each other
- However we can think of situations where the events are not really independent. For example, if we are interested in the probability of a certain number of weddings over one summer, we should take into consideration that weekends are more popular for weddings.
- In this case, a Poisson model may sometimes still be reasonable if we allow it to have a different rate for different times; we could model the rate as higher on weekends than on weekdays.
- The idea of modeling rates for a Poisson distribution against a second variable (day of the week) forms the foundation of some more advanced methods called *generalized linear models*.

Exponential Distribution

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Let $\lambda > 0$. A random variable *X* follows the exponential distribution with parameter λ if it has a continuous distribution with pf:

$$f(x \mid \lambda) = \begin{cases} \lambda e^{-\lambda x} & x > 0\\ 0 & \text{otherwise} \end{cases}$$

Parameter space: $\lambda \in [0, \infty)$. $E(X) = \frac{1}{\lambda}, \quad Var(X) = \frac{1}{\lambda^2}$

Exponential Distribution

Relationship of the Poisson and Exponential distribution

X: Time between two consecutive arrivals N_t : Number of arrivals during time *t*

$$P(X > x) = P(N_t = N_{t+x}) = P(N_x = 0) = \frac{\lambda^0 e^{-\lambda}}{0!} = e^{-\lambda}$$
$$P(X \le x) = 1 - P(X < x) = 1 - e^{-\lambda}$$

Memorylessness of the Exponential

 $P(X > x + a \mid X > a)$

$$= \frac{P(X > x + a, X > a)}{P(X > a)}$$
$$= \frac{P(X > x + a)}{P(X > a)}$$
$$= \frac{1 - F_X(x + a)}{1 - F_X(a)}$$
$$= \frac{e^{-\lambda(x + a)}}{e^{-\lambda a}}$$
$$= e^{-\lambda x}$$
$$= P(X > x).$$

Poisson Example

The number of emails that I get in a weekday can be modelled by a Poisson distribution with an average of 0.2 emails per minute.

- 1. What is the probability that I get no emails in an interval of length 5 minutes?
- 2. What is the probability that I get more than 3 emails in an interval of length 10 minutes?
- **3.** I just got an email. What is the probability that I will wait more than 3 minutes until the next email?