

# Probabilistic Graphical Models

Frequentist Estimation,  
Bayesian Networks

# Frequentist Estimation

$\theta$  is an unknown number

Each  $\theta$  defines a different probabilistic model for your data

Some models explain your data better than others

$P(x_1, \dots, x_n; \theta)$  is the likelihood function (not a conditional probability since  $\theta$  is not random)

Estimation: Find  $\hat{\theta}$

The estimator is a random variable (why?)

# Frequentist Estimation

Given  $X = x$ , the maximum likelihood estimate (MLE) will be a function of  $x$ .

Notation:  $\hat{\theta} = \delta(X_1, \dots, \dots X_n)$

Potentially confusing notation: Sometimes  $\hat{\theta}$  is used for both the estimator and the estimate.

- Note: The MLE is required to be in the parameter space  $\Omega$ .
- Often it is easier to maximize the log-likelihood  $L(\theta) = \log(f(x | \theta))$

# Maximum Likelihood Estimation

- Let  $X \sim \text{Binomial}(\theta)$ . Find the maximum likelihood estimator of  $\theta$ . Say we observe  $X = 3$ , what is the maximum likelihood estimate of  $\theta$  ?
- Let  $X_1, \dots, X_n$  be i.i.d.  $N(\mu, \sigma^2)$ .
- Find the MLE of  $\mu$  when  $\sigma^2$  is known.
- Find the MLE of  $\mu$  and  $\sigma^2$  (both unknown).
- Let  $X_1, \dots, X_n$  be i.i.d. Uniform  $[0, \theta]$ , where  $\theta > 0$ . Find  $\hat{\theta}$
- Let  $X_1, \dots, X_n$  be i.i.d. Uniform  $[\theta, \theta + 1]$ . Find  $\hat{\theta}$

# Quantifying Uncertainty

- Suppose  $X = (X_1, \dots, X_n)$  is a random sample from  $f(x | \theta)$ .
- A function  $r(X_1, \dots, X_n)$  is a statistic (and a random variable).
- A sampling distribution: the distribution of a statistic (given  $\theta$ )
- Estimator  $\hat{\theta}$  is a statistic
- Can use the sampling distributions to compare different estimators and quantify uncertainty
- Can be used to estimate number of samples we need to limit bias
- Leads to definitions of new distributions, e.g.,  $\chi_m^2$  and  $t_m$ .

# Quantifying Uncertainty

Let  $X_1, \dots, X_n$  be a random sample from a  $\mathcal{N}(\mu, \sigma^2)$  with unknown  $\mu, \sigma^2$ .  
The sample mean and the sample variance are defined as

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i, \quad \hat{\sigma}_{MLE}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2, \quad S_n = \left( \frac{\sum (X_i - \bar{X}_n)^2}{n-1} \right)^{1/2}$$

If you know  $\mu$  but not  $\sigma^2$

$$\frac{n\hat{\sigma}_{MLE}^2}{\sigma^2} \sim \chi_n^2$$

If you do not know  $\mu$  or  $\sigma^2$ , then

$$n^{1/2}(\bar{X}_n - \mu)/S_n \sim t_{n-1}$$

# Confidence interval

You can compute

$$P \left( \bar{X}_n - \frac{cS_n}{n^{1/2}} < \mu < \bar{X}_n + \frac{cS_n}{n^{1/2}} \right) \geq \gamma$$

- $\gamma$  – Confidence Interval for  $\mu$ .
- $\mu$  is not random, the interval is.
- Interpretation:  $\gamma$  is the frequency we expect the random interval to include the true value, if we repeat the experiment multiple times

# Properties of an Estimator

An estimator  $\hat{\theta} = g(X_1, \dots, X_n)$  is a function of random variables  $X_1, \dots, X_n$  and therefore has a distribution. The distribution of  $\hat{\theta}$  is called sampling distribution.

## Unbiased estimator

An estimator is unbiased if  $E(\hat{\theta}) = \theta$ .  $E(\hat{\theta}) - \theta$  is called the bias of the estimator.

## Consistent estimator

An estimator is consistent if  $\hat{\theta}_n \xrightarrow{p} \theta$ .

## Example

Bernoulli MLE  $\hat{\theta}_{MLE} = \sum_{i=1}^n x_i$ . Is it unbiased? Is it consistent?



# Mean Squared Error of an Estimator

## Mean squared error estimator

$$\begin{aligned} \text{MSE}(\hat{\theta}) &= \mathbb{E}[(\hat{\theta} - \theta)^2] \\ &= \text{Var}(\hat{\theta} - \theta) + (\mathbb{E}[\hat{\theta} - \theta])^2 \\ &= \text{Var}(\hat{\theta}) + \text{Bias}^2(\hat{\theta}) \end{aligned}$$

For unbiased estimators,  $\text{MSE}(\hat{\theta}) = \text{Var}(\hat{\theta})$

# Sufficient Statistics

- A statistic:  $T = r(X_1, \dots, X_n)$

Def: Sufficient Statistics

Let  $X_1, \dots, X_n$  be a random sample from  $f(x | \theta)$  and let  $T$  be a statistic.  
If the conditional distribution of

$$X_1, \dots, X_n | T = t$$

does not depend on  $\theta$  then  $T$  is called a sufficient statistic.

- The idea: Just as good to have the observed sufficient statistic as it is to have the individual observations of  $X_1, \dots, X_n$ .
- Can limit our search for a good estimator to sufficient statistics

# Sufficient Statistics

- Theorem: Factorization Criterion

Let  $X_1, \dots, X_n$  be a random sample from  $f(x | \theta)$  where  $\theta \in \Omega$  is unknown. A statistic  $T = r(X_1, \dots, X_n)$  is a sufficient statistic for  $\theta$  if and only if for all  $\mathbf{x} \in \mathbb{R}^n$  and all  $\theta \in \Omega$ , the joint pdf/pf  $f_n(\mathbf{x} | \theta)$  can be factored as

$$f_n(\mathbf{x} | \theta) = u(\mathbf{x})v(r(\mathbf{x}), \theta)$$

where function  $u$  and  $v$  are nonnegative.

- The function  $u$  may depend on  $\mathbf{x}$  but not on  $\theta$
- The function  $v$  depends on  $\theta$  but depends on  $\mathbf{x}$  only through the value of the statistic  $r(\mathbf{x})$

Both MLEs and Bayesian estimators depend on data only through sufficient statistics.

# MLE vs Bayesian Estimation

## MLE

Does not always exist

Is not always appropriate

Is not always unique

## Bayes

More difficult computationally

Not a single point

# Probabilistic Graphical Models

## Directed graphical models

- Bayes Nets
- Conditional dependence

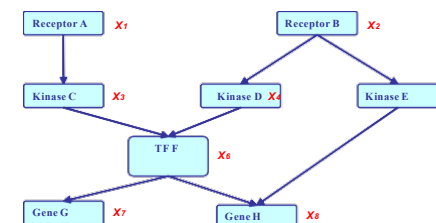
## Undirected graphical models

- Markov random fields (MRFs)
- Factor graphs

# Two types of GMs

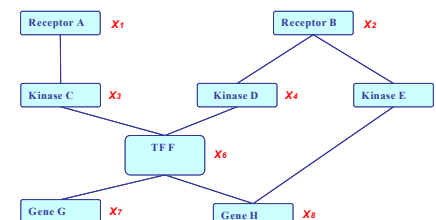
- **Directed edges** give **causality** relationships (Bayesian Network or Directed Graphical Model):

$$\begin{aligned} &P(X_1, X_2, X_3, X_4, X_5, X_6, X_7, X_8) \\ &= P(X_1) P(X_2) P(X_3|X_1) P(X_4|X_2) P(X_5|X_2) \\ &\quad P(X_6|X_3, X_4) P(X_7|X_6) P(X_8|X_5, X_6) \end{aligned}$$



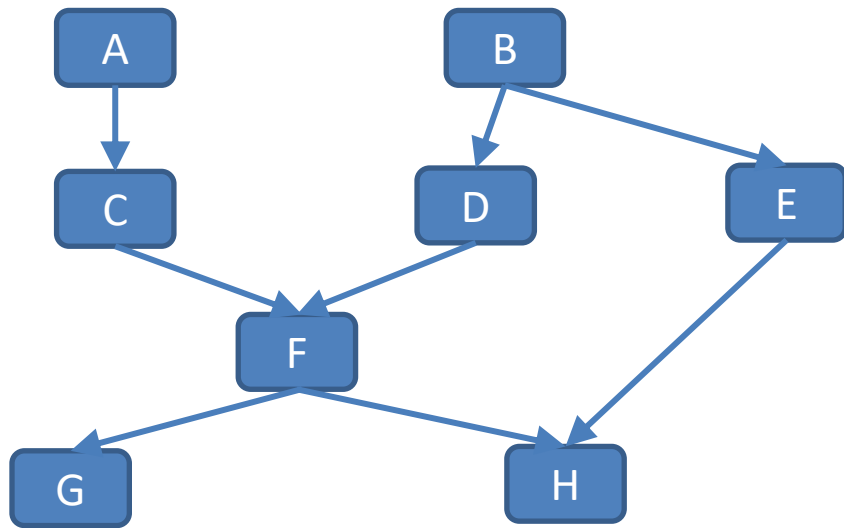
- **Undirected edges** simply give **correlations** between variables (Markov Random Field or Undirected Graphical model):

$$\begin{aligned} &P(X_1, X_2, X_3, X_4, X_5, X_6, X_7, X_8) \\ &= \frac{1}{Z} \exp\{E(X_1) + E(X_2) + E(X_3, X_1) + E(X_4, X_2) + E(X_5, X_2) \\ &\quad + E(X_6, X_3, X_4) + E(X_7, X_6) + E(X_8, X_5, X_6)\} \end{aligned}$$



# Directed Graphical Models

A Directed Acyclic Graph



A joint Probability Distribution

$$P(A, B, C, D, E, F, G, H)$$

$$P(A, \dots, H)$$

$$= \prod_{V \in \{A, \dots, H\}} P(V | Pa_G(V))$$

(Local) Markov Condition:

Every variable is independent of its non-descendants given its parents (in the graph)

# Example: Expert systems

- Beinlich et al. 1989
- Encodes medical knowledge
- Patient monitoring system
  - Measurements:
    - Blood pressure 120/80 mmHg
    - Heart rate 80/min
    - Respiratory rate 10/min
    - ...
  - Query:
    - $\Pr(\text{kinked tube}=\text{true} \mid \text{measurements}) = ?$

## The ALARM Monitoring System: A Case Study with two Probabilistic Inference Techniques for Belief Networks

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**Abstract** ALARM (A Logical Alarm Reduction Mechanism) is a diagnostic application used to explore probabilistic reasoning techniques in belief networks. ALARM implements an alarm message system for patient monitoring; it calculates probabilities for a differential diagnosis based on available evidence. The medical knowledge is encoded in a graphical structure connecting 8 diagnoses, 16 findings and 13 intermediate variables. Two algorithms were applied to this belief network: (1) a message-passing algorithm by Pearl for probability updating in multiply connected networks using the method of conditioning; and (2) the Lauritzen-Spiegelhalter algorithm for local probability computations on graphical structures. The characteristics of both algorithms are analyzed and their specific applications and time complexities are shown.

### Introduction

The goal of the ALARM monitoring system is to provide specific text messages advising the user of possible problems. This is a diagnostic task, and we have chosen to represent the relevant knowledge in the language of a belief network (Fig.1). This graphical representation [Pearl 86b] facilitates the integration of qualitative and quantitative knowledge, the assessment of multiple faults, as required by our domain, and nonmonotonic and bidirectional reasoning.

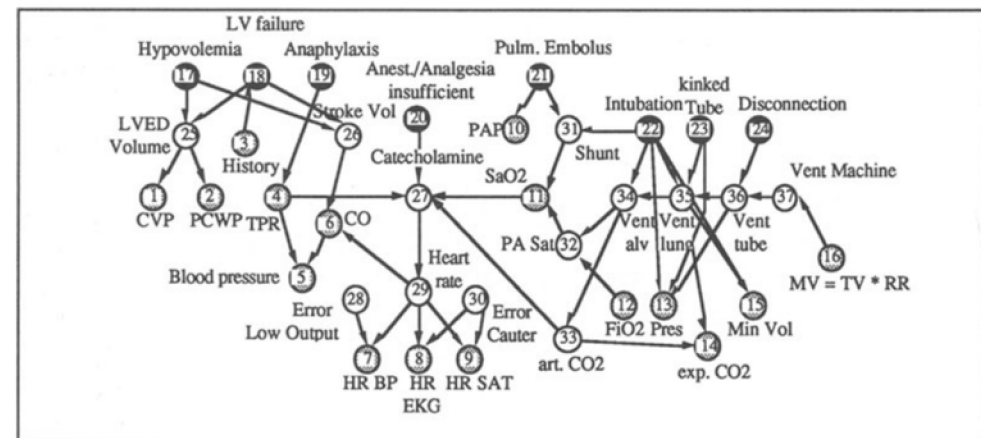
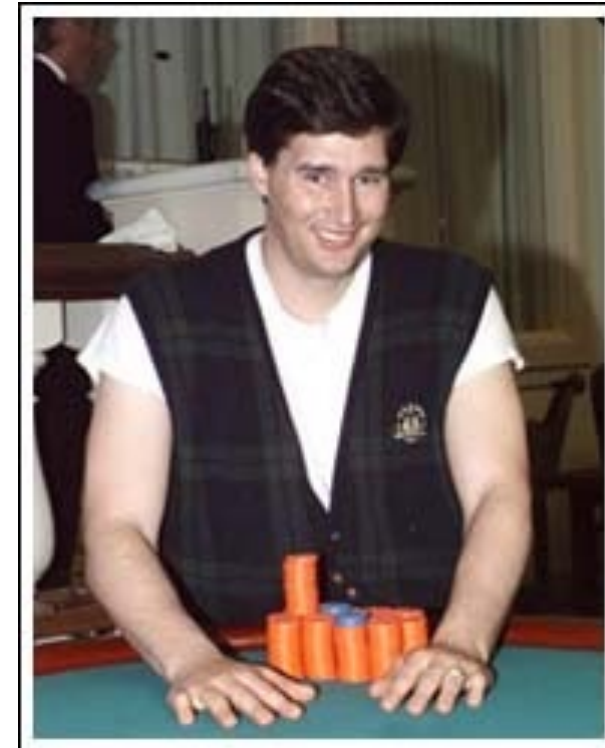
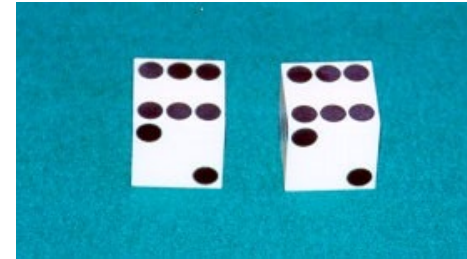


Fig. 1 The ALARM network representing causal relationships is shown with diagnostic (●), intermediate (○) and measurement (⊙) nodes. CO: cardiac output, CVP: central venous pressure, LVED volume: left ventricular end-diastolic volume, LV failure: left ventricular failure, MV: minute ventilation, PA Sat: pulmonary artery oxygen saturation, PAP: pulmonary artery pressure, PCWP: pulmonary capillary wedge pressure, Pres: breathing pressure, RR: respiratory rate, TPR: total peripheral resistance, TV: tidal volume



# Example: The Dishonest Casino

- A casino has two dice:
  - Fair die
    - $P(1) = P(2) = P(3) = P(5) = P(6) = 1/6$
  - Loaded die
    - $P(1) = P(2) = P(3) = P(5) = 1/10$
    - $P(6) = 1/2$
  - Casino player switches back-&-forth between fair and loaded die once every 20 turns
- Game:
  - You bet \$1
  - You roll (always with a fair die)
  - Casino player rolls (maybe with fair die, maybe with loaded die)
  - Highest number wins \$2



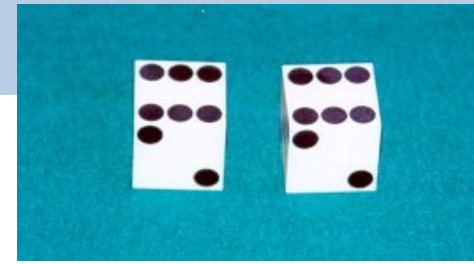
# Puzzles regarding the dishonest casino

**GIVEN:** A sequence of rolls by the casino player

1245526462146146136136661664661636616366163616515615115146123562344

## QUESTION

- How likely is this sequence, given our model of how the casino works?
  - This is the **EVALUATION** problem
- What portion of the sequence was generated with the fair die, and what portion with the loaded die?
  - This is the **DECODING** question
- How “loaded” is the loaded die? How “fair” is the fair die? How often does the casino player change from fair to loaded, and back?
  - This is the **LEARNING** question



- ❑ Picking variables
  - ❑ Observed
  - ❑ Hidden
  - ❑ Discrete
  - ❑ Continuous
- ❑ Picking structure
  - ❑ CAUSAL
  - ❑ Generative
  - ❑ Coupling
- ❑ Picking Probabilities
  - ❑ “Natural”
  - ❑ Zero probabilities
  - ❑ Orders of magnitudes
  - ❑ Relative values

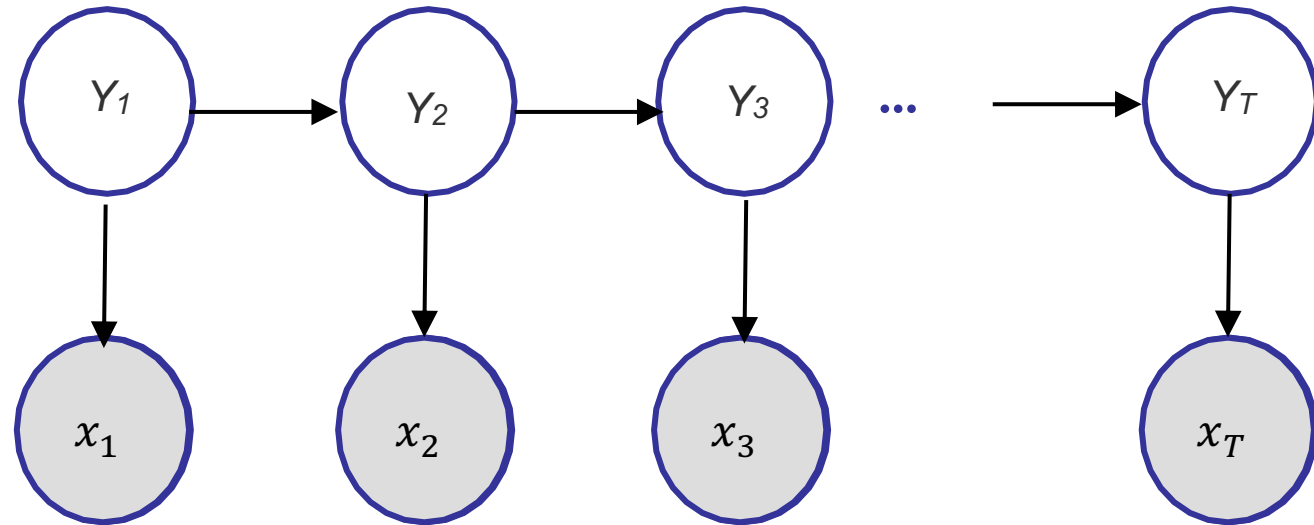
# Hidden Markov Model

## The underlying source:

Speech signal  
genome function  
dice

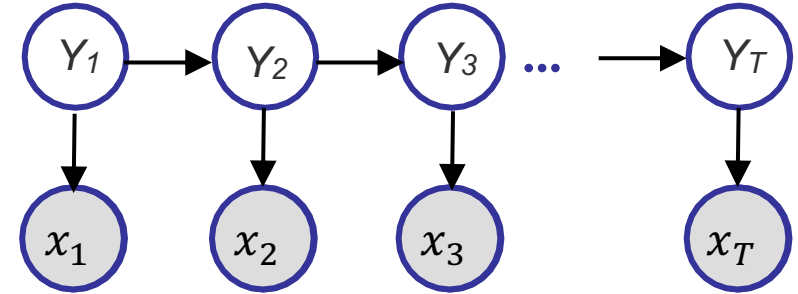
## The sequence:

Phonemes  
DNA sequence  
sequence of rolls



# Probability of a parse

- Given a sequence  $\mathbf{x} = x_1, \dots, x_T$  and a parse  $\mathbf{y} = y_1, \dots, y_T$ ,
- To find how likely is the parse: (given our HMM and the sequence)



$$\begin{aligned} p(\mathbf{x}, \mathbf{y}) &= p(x_1, \dots, x_T, y_1, \dots, y_T) && \text{(Joint probability)} \\ &= p(y_1) p(x_1 | y_1) p(y_2 | y_1) p(x_2 | y_2) \dots p(y_T | y_{T-1}) p(x_T | y_T) \\ &= p(y_1) P(y_2 | y_1) \dots p(y_T | y_{T-1}) \times p(x_1 | y_1) p(x_2 | y_2) \dots p(x_T | y_T) \\ &= p(y_1, \dots, y_T) p(x_1, \dots, x_T | y_1, \dots, y_T) \end{aligned}$$

- Marginal probability:  $p(\mathbf{x}) = \sum_{\mathbf{y}} p(\mathbf{x}, \mathbf{y}) = \sum_{\mathbf{y}}$
- Posterior probability:  $p(\mathbf{y} | \mathbf{x}) = p(\mathbf{x}, \mathbf{y}) / p(\mathbf{x})$
- We will learn how to do this efficiently (polynomial time)

# Bayesian Network

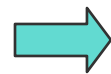
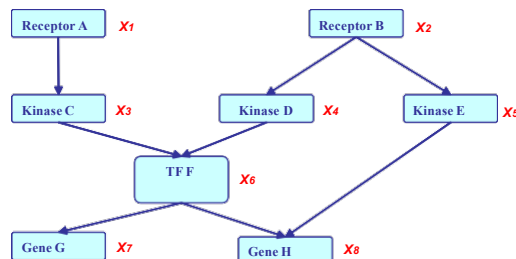
- A BN is a directed graph whose nodes represent the random variables and whose edges represent direct influence of one variable on another.
- It is a data structure that provides the skeleton for representing **a joint distribution** compactly in a **factorized** way;
- It offers a compact representation for **a set of conditional independence assumptions** about a distribution;
- We can view the graph as encoding a **generative sampling process** executed by nature, where the value for each variable is selected by nature using a distribution that depends only on its parents. In other words, each variable is a stochastic function of its parents.

# Bayesian Network: Factorization

Given a DAG, The most general form of the probability distribution that is **consistent with** the graph factors according to “node given its parents”:

$$P(\mathbf{X}) = \prod_{i=1:d} P(X_i | \mathbf{X}_{\pi_i})$$

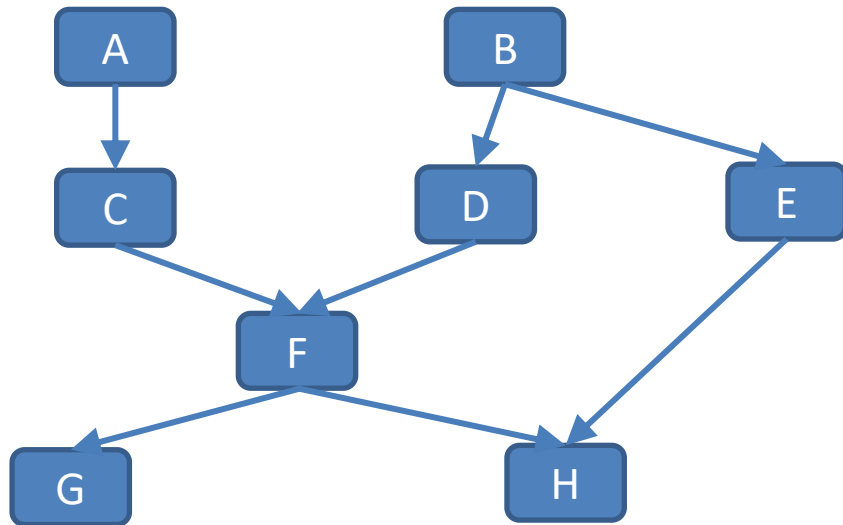
where  $\mathbf{X}_{\pi_i}$  is the set of parents of  $X_i$ ,  $d$  is the number of nodes (variables) in the graph.



$$\begin{aligned} &P(X_1, X_2, X_3, X_4, X_5, X_6, X_7, X_8) \\ &= P(X_1) P(X_2) P(X_3 | X_1) P(X_4 | X_2) P(X_5 | X_2) \\ &P(X_6 | X_3, X_4) P(X_7 | X_6) P(X_8 | X_5, X_6) \end{aligned}$$

# Specification of a directed GM

- There are two components to any GM:
  - the *qualitative* specification (graph)
  - the *quantitative* specification (jpd)



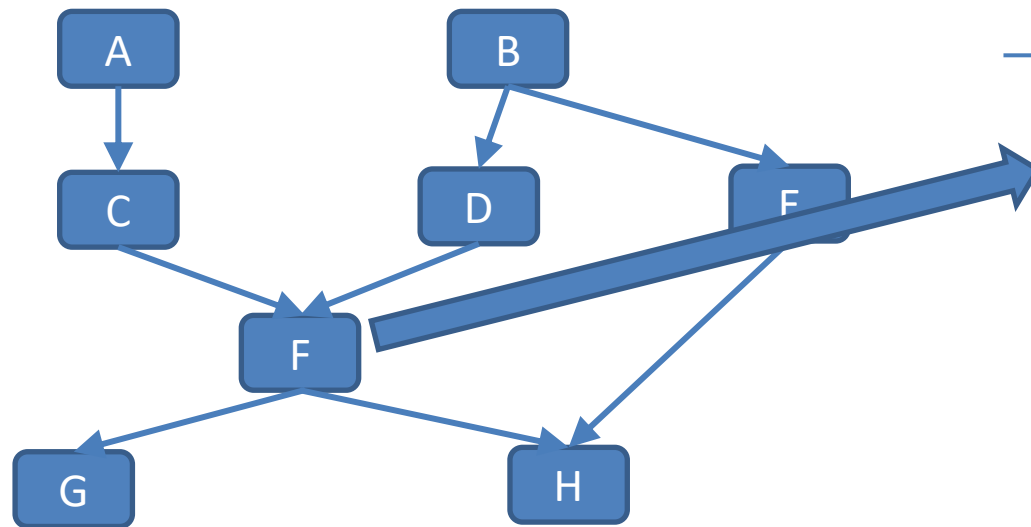
The graph dictates the factorization of the joint probability distribution

$$P(A, \dots, H) = \prod_{V \in \{A, \dots, H\}} P(V | Pa_G(V))$$



# Specification of a directed GM

- There are two components to any GM:
  - the *qualitative* specification (graph)
  - the *quantitative* specification (jpd)



<i>C</i>	<i>D</i>	$P(F   C, D)$	
0	0	0.9	0.1
1	0	0.2	0.8
0	1	0.9	0.1
1	1	0.01	0.99

# Qualitative Specification

- Where does the qualitative specification come from?
  - Prior knowledge of causal relationships
  - Prior knowledge of modular relationships
  - Assessment from experts
  - Learning from data
  - We simply like a certain architecture (e.g. a layered graph)
  - ...

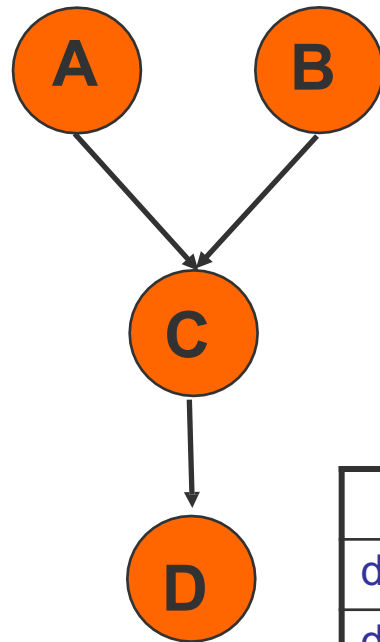
# Quantitative Specification

## Conditional probability tables (CPTs)

$a^0$	0.75
$a^1$	0.25

$b^0$	0.33
$b^1$	0.67

$$P(a,b,c,d) = P(a)P(b)P(c|a,b)P(d|c)$$



	$a^0b^0$	$a^0b^1$	$a^1b^0$	$a^1b^1$
$c^0$	0.45	1	0.9	0.7
$c^1$	0.55	0	0.1	0.3

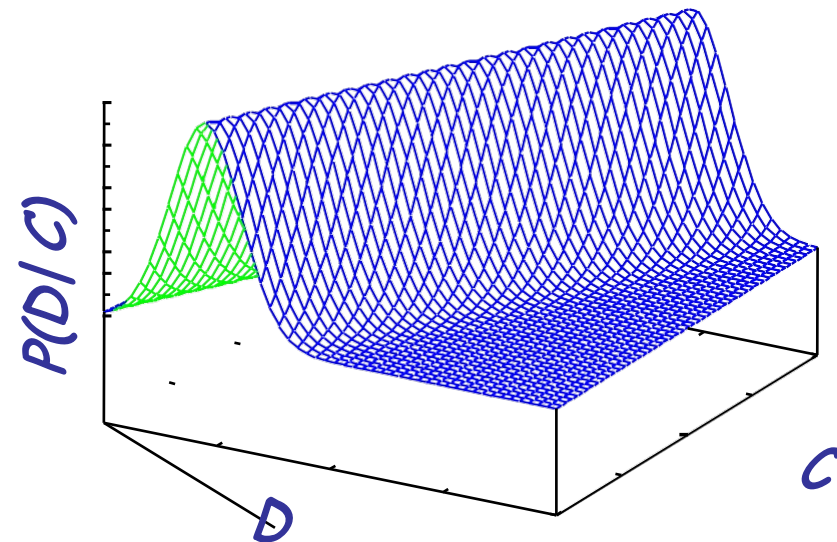
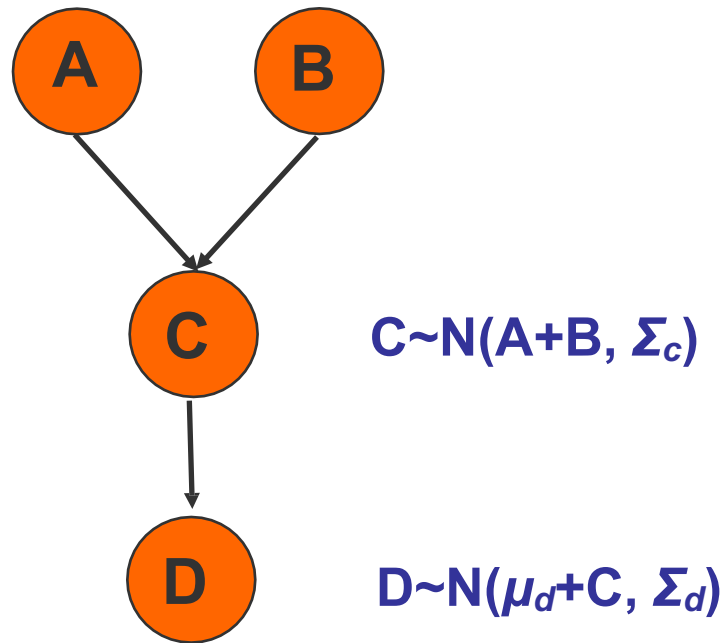
	$c^0$	$c^1$
$d^0$	0.3	0.5
$d^1$	0.7	0.5

# Qualitative Specification

## Conditional probability density func. (CPDs)

$$A \sim N(\mu_a, \Sigma_a) \quad B \sim N(\mu_b, \Sigma_b)$$

$$P(a,b,c,d) = P(a)P(b)P(c|a,b)P(d|c)$$

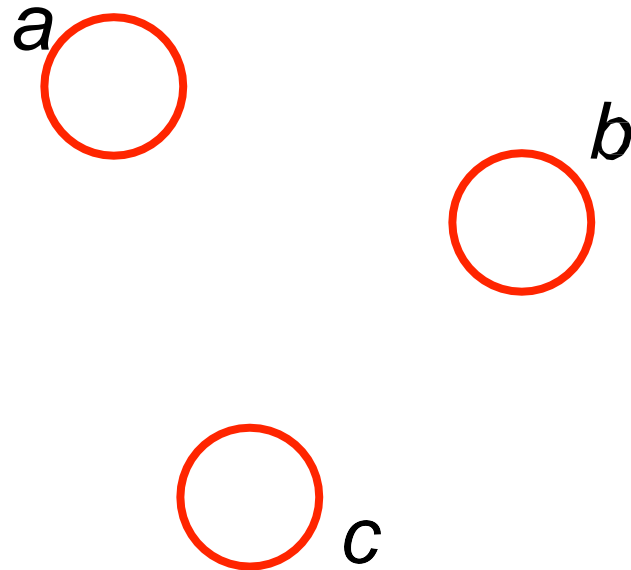


# Qualitative Specification

- Where does the qualitative specification come from?
  - Prior knowledge of causal relationships
  - Prior knowledge of modular relationships
  - Assessment from experts
  - Learning from data
  - We simply like a certain architecture (e.g. a layered graph)
  - ...
- Graphs imply some conditional independencies. (wherever you got them)
  - What does this mean?
  - For every distribution that factorizes according to the graph

# Implied Independencies.

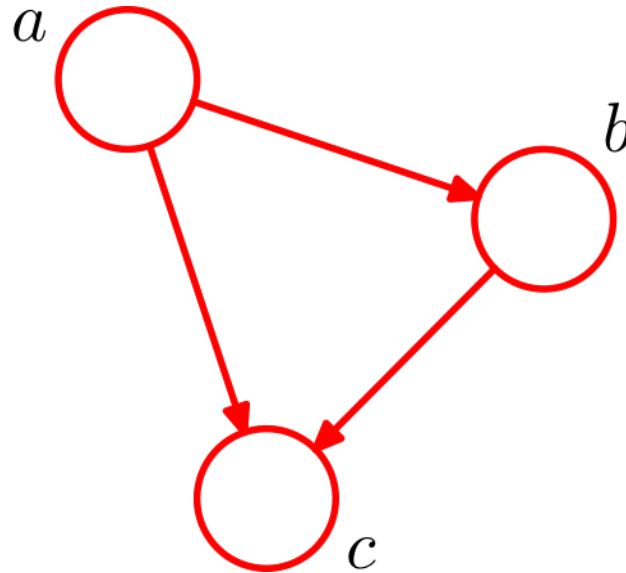
*Are  $a$  and  $b$  independent (  $a \perp b$  )?*



$$\mathbf{p(a,b,c) = p(a)p(b)p(c)}$$

# Implied Independencies

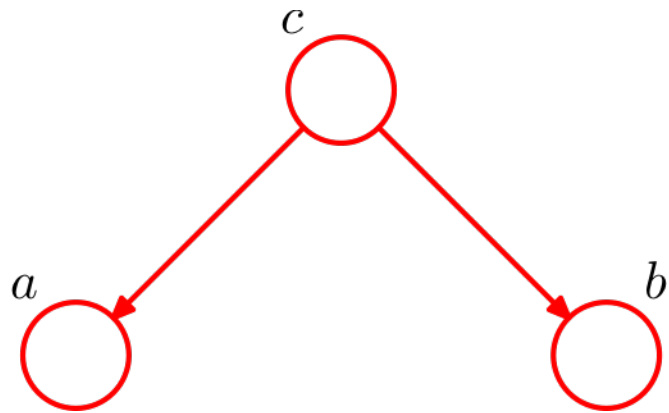
$$p(a,b,c) = p(a)p(b|a)p(c|a,b)$$



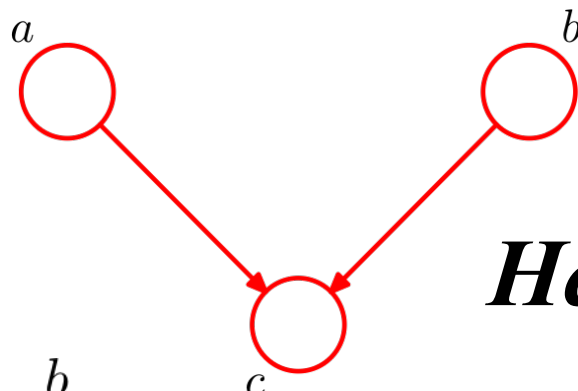
Note there are **no conditional independencies** (fully connected graph)

# Three interesting cases

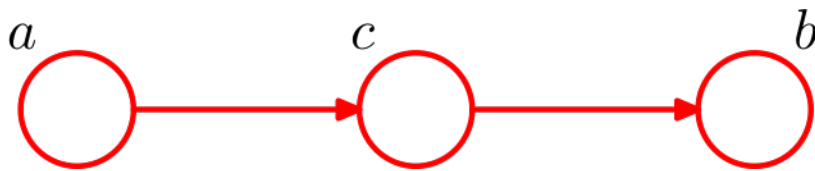
*Tail-to-tail*



*Head-to-head*

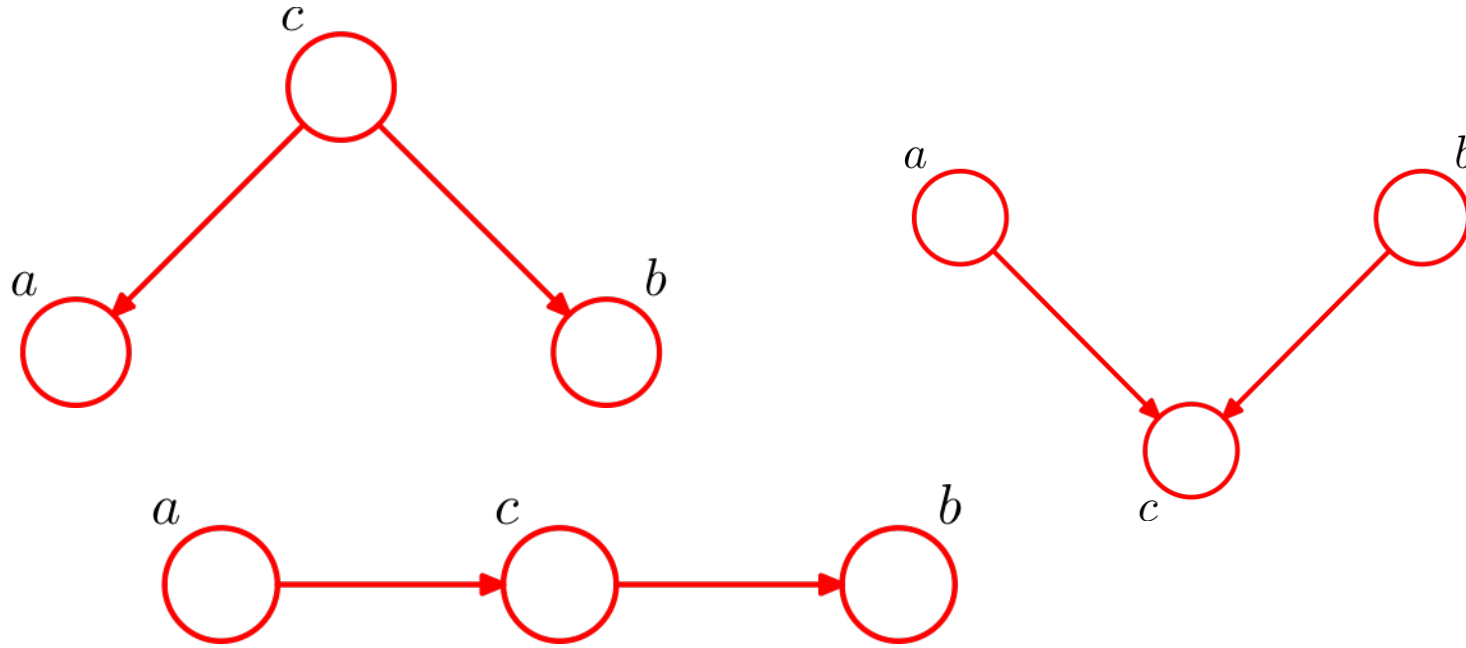


*Head-to-tail*





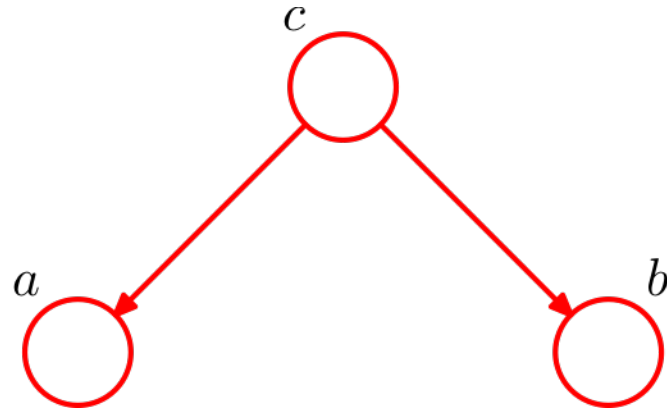
# Three interesting cases



For each case, consider two questions:

- 1) Is  $a \perp b$  ?
- 2) Is  $a \perp b \mid c$  ? (i.e.  $c$  is observed)

# Case one (tail-to-tail)

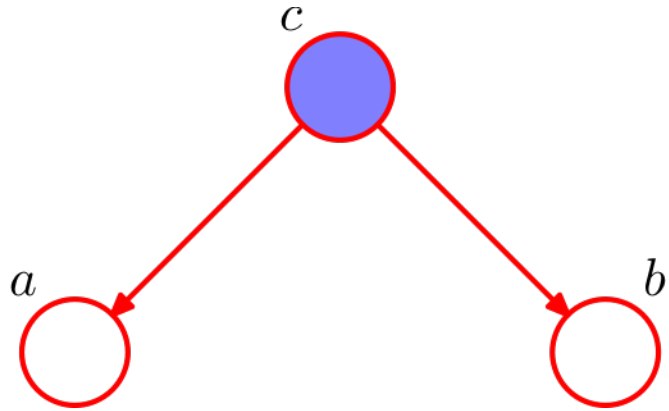


$a \not\perp b$

This graph represents  $P(a, b, c) = P(c)P(a|c)P(b|c)$

To prove independence, we need to come up with a counter-example

# Case one (tail-to-tail)



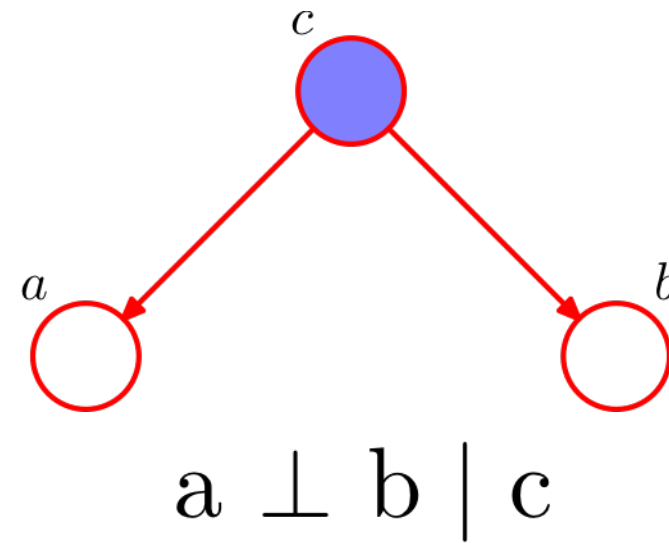
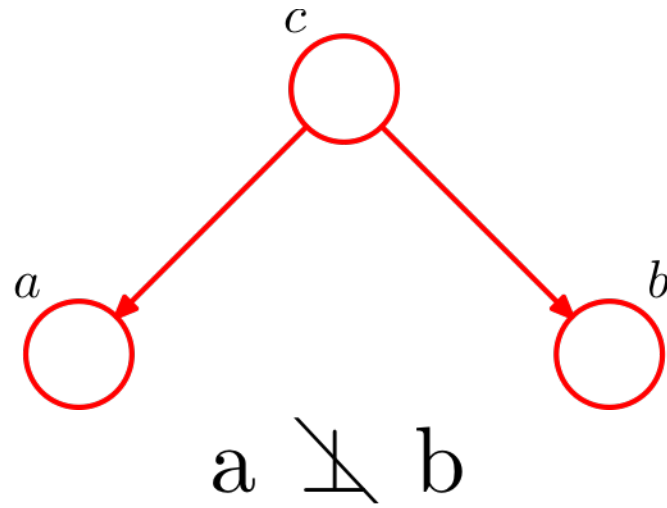
$$a \perp b \mid c$$

$$p(a, b, c) = p(c)p(a|c)p(b|c) \quad (\text{what the graph represents in general})$$

$$p(a, b|c) = p(a|c)p(b|c) \quad (\text{with } c \text{ observed})$$

This is the definition of  $a \perp b \mid c$

# Case one (tail-to-tail) summary

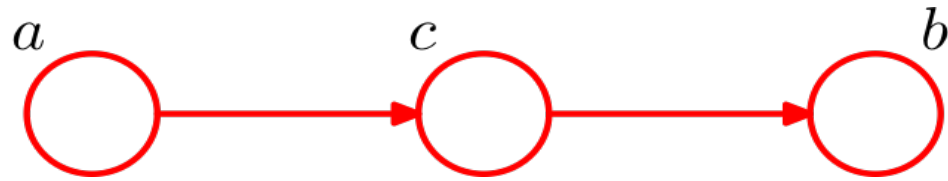


Tail-to-tail case

With no conditioning, no independence ( $\exists P$ )

With conditioning, we have independence

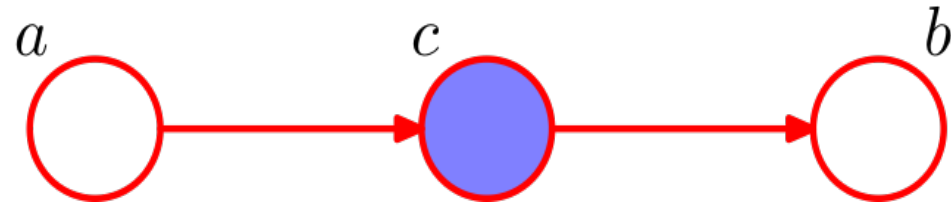
## Case two (head-to-tail)



$a \perp\!\!\!\perp b$

This graph represents  $P(a, b, c) = P(a)P(a|c)P(b|c)$

# Case two (head-to-tail)

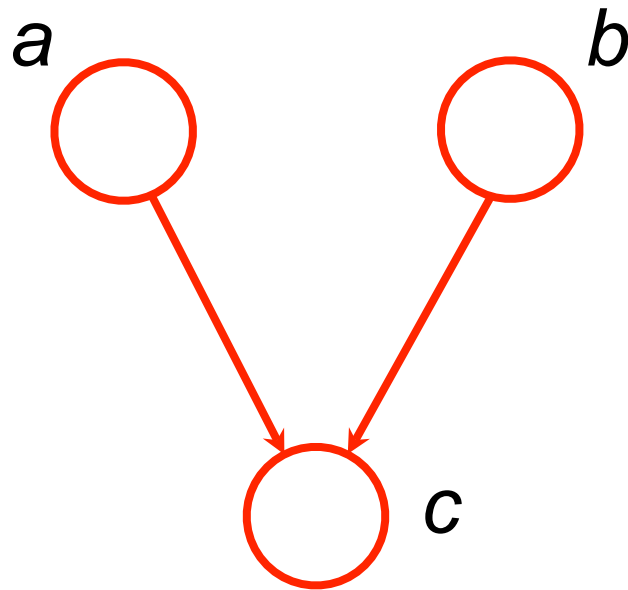


$$a \perp b \mid c$$

$$\begin{aligned} p(a, b \mid c) &= \frac{p(a, b, c)}{p(c)} && \text{(definition)} \\ &= \frac{p(a)p(c|a)p(b|c)}{p(c)} && \text{(from graph)} \\ &= \frac{p(a)p(a|c)p(c)p(b|c)}{p(a)p(c)} && \text{(Bayes on } p(c|a)) \\ &= p(a|c)p(b|c) \end{aligned}$$

## Case three (head-to-head)

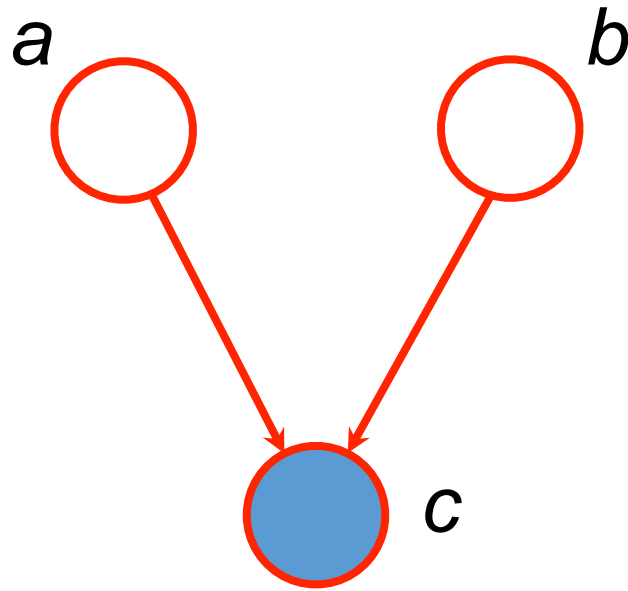
*Are  $a$  and  $b$  independent ( $a \perp b$ )?*



$$p(a, b) = \sum_c p(a)p(b)p(c | a, b) = p(a)p(b)$$

## Case three (head-to-head)

*Are  $a$  and  $b$  conditionally independent ( $a \perp b \mid c$ )?*

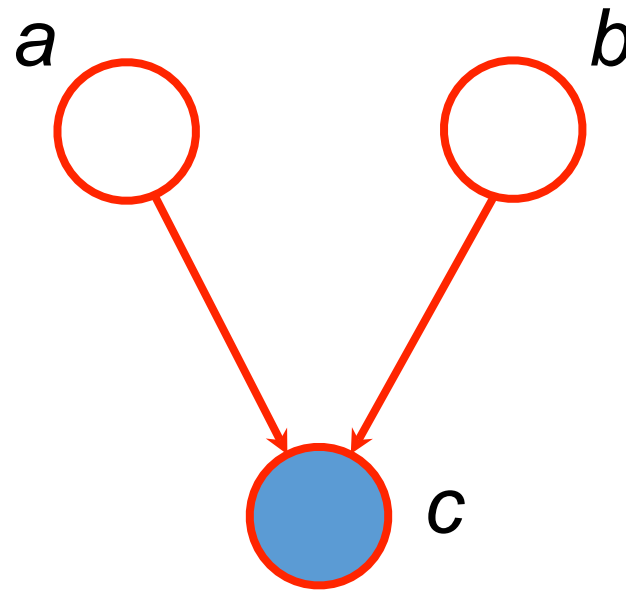


$$p(a,b,c) = p(a)p(b)p(c|a,b)$$



# Case three (head-to-head)

Are  $a$  and  $b$  conditionally independent ( $a \perp b \mid c$ )?

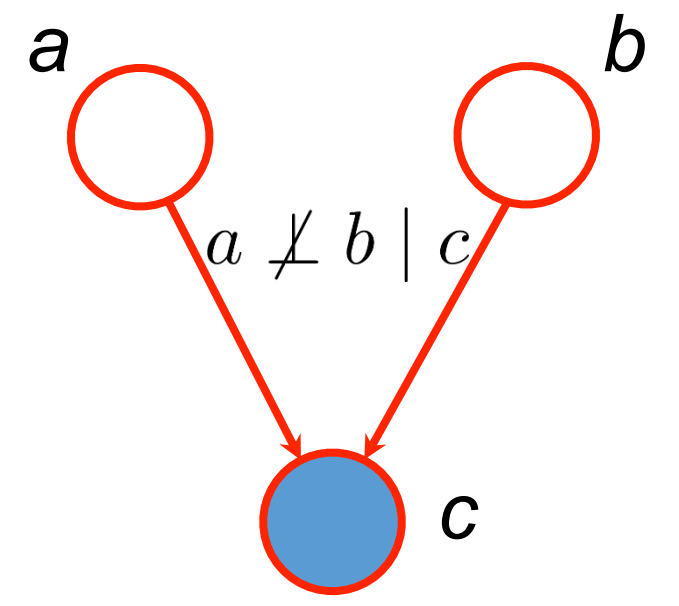
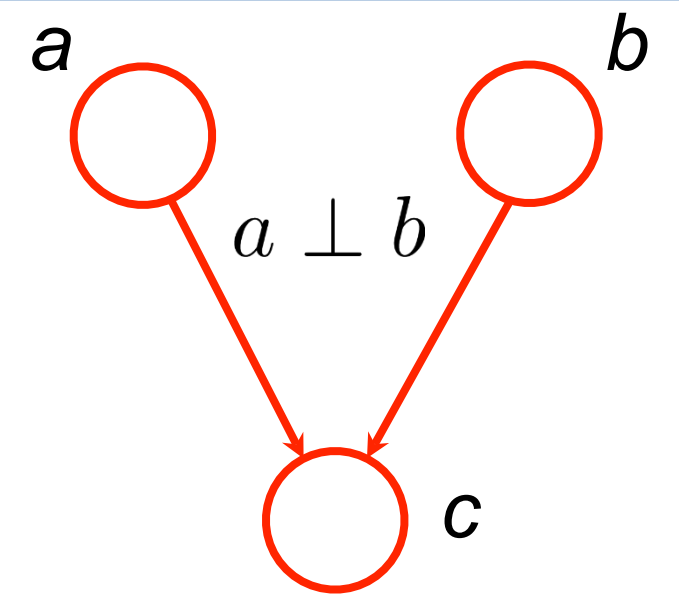
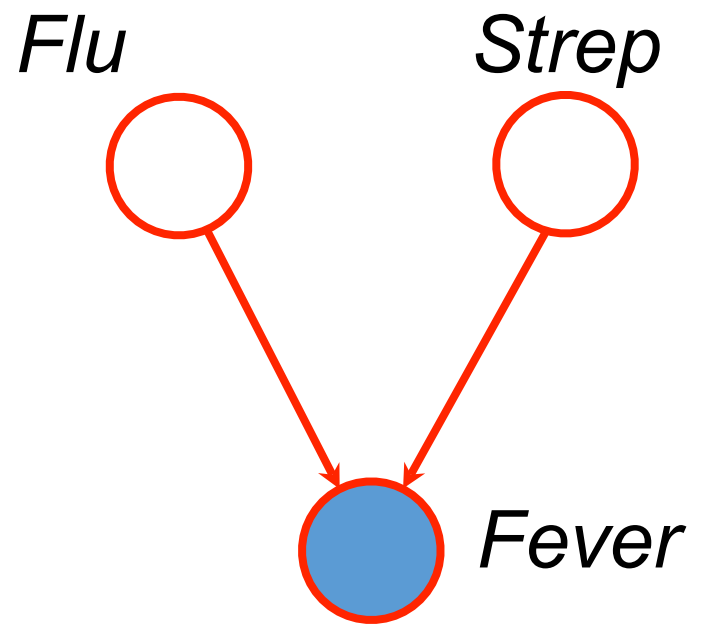


$$\begin{aligned} p(a,b|c) &= \frac{p(a,b,c)}{p(c)} \\ &= \frac{p(a)p(b)p(c|a,b)}{p(c)} \\ &\neq p(a|c)p(b|c) \quad (\text{in general}) \end{aligned}$$

Attempt at algebraic proof.

Unless the algebra reduces to something obviously false, we typically look for a counter example

# Case three (head-to-head)



Phenomenon in Bayes networks known as **explaining away**

# Summary

Bayesian networks: Graph (DAG)+JPD

JPD factorizes according to the factorization theorem

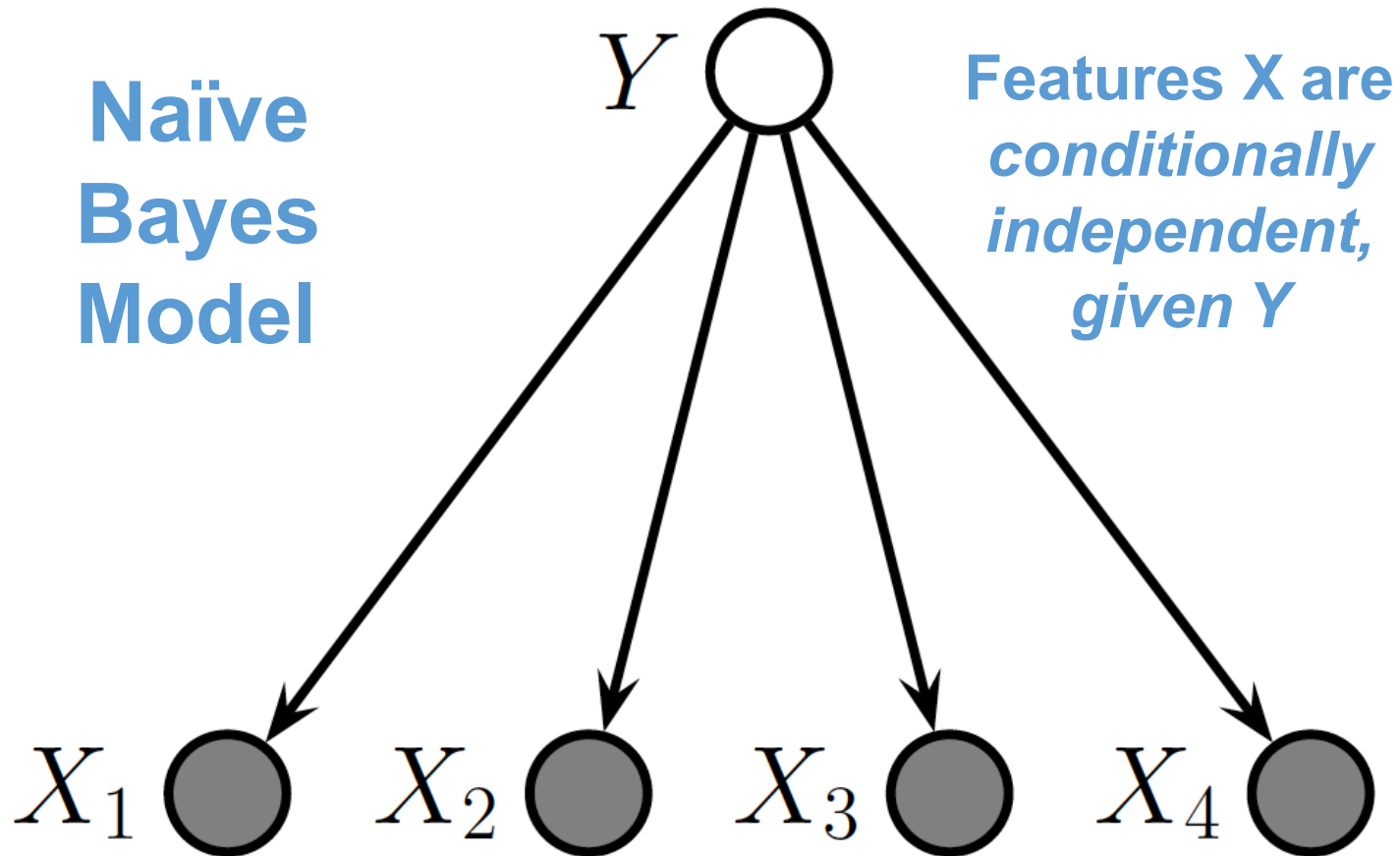
Factorization theorem implies a set of conditional independencies.

Next: A general algorithm for reading independencies from the graphs.

# Shading & Plate Notation

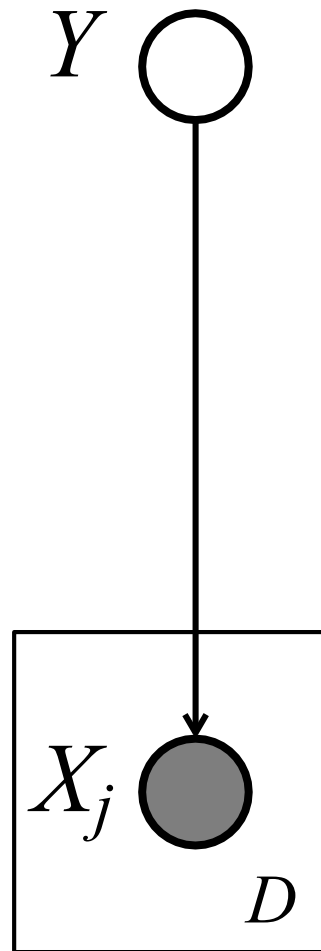
*Convention: Shaded nodes are observed, open nodes are latent/hidden/unobserved*

**Naïve  
Bayes  
Model**



**Features  $X$  are  
conditionally  
independent,  
given  $Y$**

$$p(y, \mathbf{x}) = p(y) \prod_{j=1}^D p(x_j | y)$$



*Plates denote  
replication of  
random variables*

# Example: Gaussian Mixture Model

## Probability Model

$$\pi \sim \text{Dirichlet}(\cdot)$$

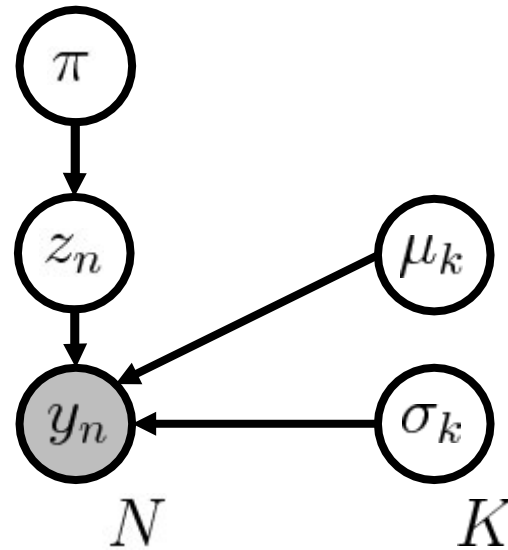
$$\mu_k \sim \mathcal{N}(\cdot)$$

$$\sigma_k \sim \text{Inv-Gamma}(\cdot)$$

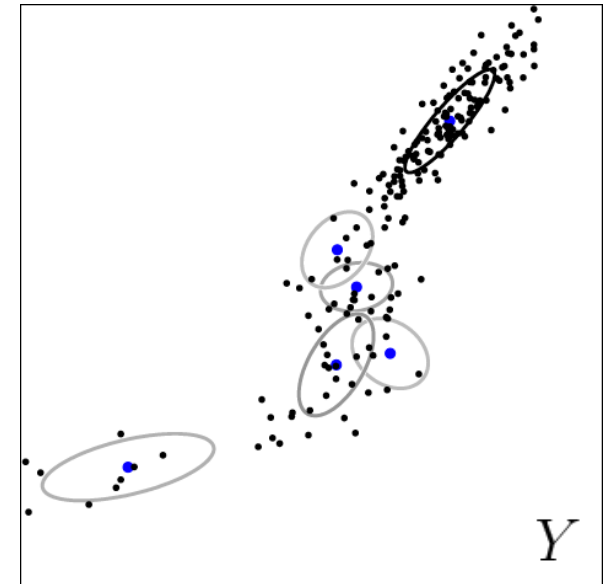
$$z_n \mid \pi \sim \text{Cat}(\pi)$$

$$y_n \mid z_n, \mu_{z_n}, \sigma_{z_n} \sim \mathcal{N}(\mu_{z_n}, \sigma_{z_n})$$

## Bayes Net



## Joint Sample



*Sample all nodes with no parents, then children, etc., to terminals. Can sample nodes at same level in parallel.*